



# TIMESPAN

Management of chronic cardiometabolic disease and treatment discontinuity in adult ADHD patients

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# Abbreviations

T2D	Type-2 Diabetes
CVD	cardiovascular disease
ADHD	Attention Deficit Hyperactivity Disorder
NOS	Newcastle-Ottawa Scale
GRACE	Good Research for Comparative Effectiveness

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## 1. Executive Summary

Cardiometabolic disease, including Obesity, Type-2 Diabetes (T2D) and cardiovascular disease (CVD) are well-known chronic conditions that frequently co-occurs. Each of these conditions are associated with premature death, as well as high levels of health care costs. Attention Deficit Hyperactivity Disorder (ADHD) is a common neurodevelopmental condition, characterized by hyperactivity, inattention, impulsivity, that often persist from childhood into adulthood. It is a costly disorder that often co-occurs with substance use disorders bipolar disorder, and depression. It is associated with criminality, occupational problems, and premature death, including suicide.

Research on psychiatric comorbidity in ADHD has informed clinical treatment guidelines to help mental health teams to build up an expertise in the diagnosis and treatment of individuals with ADHD and cooccurring psychiatric disorders. In contrast, due to a lack of research, guidelines about the management of cardiometabolic disease in individuals with ADHD is scarce.

To guide clinical practice and prevention strategies, we have conducted three systematic reviews and meta-analysis:

- i) Li L, Yao H, Zhang L, Garcia-Argibay M, Du Rietz E, Brikell I, Solmi M, Cortese S, Ramos-Quiroga JA, Ribasés M, Chang Z, and Larsson H (In press). Attention-deficit/hyperactivity disorder is associated with increased risk of cardiovascular diseases: a systematic review and meta-analysis. JCPP Advances.
- ii) Zhang L, Yao H, Li L, Du Rietz E, Andell P, Garcia-Argibay M, D'Onofrio BM, Cortese S, Larsson H, Chang Z. Risk of Cardiovascular Diseases Associated With Medications Used in Attention-Deficit/Hyperactivity Disorder: A Systematic Review and Meta-analysis. JAMA Netw Open. 2022 Nov 1;5(11):e2243597.
- Garcia-Argibay M, Li L, Du Rietz E, Zhang L, Yao H, Jendle J, Ramos-Quiroga JA, Ribasés M, Chang Z, Brikell I, Cortese S, Larsson H. The association between type 2 diabetes and attention- deficit/hyperactivity disorder: A systematic review, meta-analysis, and population-based sibling study. Neurosci Biobehav Rev. 2023 Apr;147:105076.

# 2. Deliverable report

To guide clinical practice and prevention strategies, we have conducted three systematic reviews and meta-analysis:

# i) ADHD as a risk factor for CVD

To systematically review, quantitatively synthesize, and appraise available evidence on the link between ADHD with CVDs, we searched relevant articles in PubMed, Embase, PsycINFO, and Web of Science from inception to May 1, 2022. Study quality was assessed by using the Newcastle-Ottawa Scale (NOS), and random-effects model meta-analyses were performed. A total of 18,391,169 (ADHD: n=421,224) individuals from 11 studies were included in our systematic review and 8,196,648 (ADHD=332,619) individuals from five studies were included in the main meta-analysis of adjusted estimates. Pooled estimates showed that ADHD was significantly associated with an increased risk of CVDs in analyses based on adjusted effect size (OR=1.96; 95% CI =1.19-2.23, Q=140.74,  $P_Q$ <0.001,  $I^2$ =97.2%). When restricted among adults, the heterogeneity declined to null (OR=1.73; 95% CI =1.14-2.62, Q=6.28,  $P_Q$ =0.10,  $I^2$ =6.28%), suggesting age might be the main source of heterogeneity. In

subgroup analyses, we found increased risk of CVDs associated with ADHD across age groups, type of CVDs, and data sources. This systematic review and meta-analyses indicate that ADHD is associated with increased risk for CVDs, but further studies with various study designs are warranted to advance the understanding of the underlying mechanisms for the observed association between ADHD and CVDs. Additional research is also needed to resolve the role of ADHD medications which remains unclear due to the limited number of primary studies exploring this issue.

Reference: Li L, Yao H, Zhang L, Garcia-Argibay M, Du Rietz E, Brikell I, Solmi M, Cortese S, Ramos-Quiroga JA, Ribasés M, Chang Z, and Larsson H (In press). Attention-deficit/hyperactivity disorder is associated with increased risk of cardiovascular diseases: a systematic review and meta-analysis. JCPP Advances.

# ii) ADHD medication and the risk of CVD

To provide an updated synthesis of evidence on whether ADHD medications are associated with the risk of a broad range of CVDs, we searched relevant articles in PubMed, Embase, PsycINFO, and Web of Science up to May 1, 2022. We included observational studies investigating the association between ADHD medications (including stimulants and nonstimulants) and risk of CVD. Independent reviewers extracted data and assessed study quality using the Good Research for Comparative Effectiveness (GRACE) checklist. Nineteen studies (with 3 931 532 participants including children, adolescents, and adults; 60.9% male), of which 14 were cohort studies, from 6 countries or regions were included in the meta-analysis. Median follow-up time ranged from 0.25 to 9.5 years (median, 1.5 years). Pooled adjusted relative risk (RR) did not show a statistically significant association between ADHD medication use and any CVD among children and adolescents (RR, 1.18; 95%CI, 0.91-1.53), young or middle-aged adults (RR, 1.04; 95%CI, 0.43-2.48), or older adults (RR, 1.59; 95%CI, 0.62-4.05). No significant associations for stimulants (RR, 1.24; 95%CI, 0.84-1.83) or nonstimulants (RR, 1.22; 95% CI, 0.25-5.97) were observed. For specific cardiovascular outcomes, no statistically significant association was found in relation to cardiac arrest or arrhythmias (RR, 1.60; 95%CI, 0.94-2.72), cerebrovascular diseases (RR, 0.91; 95%CI, 0.72-1.15), or myocardial infarction (RR, 1.06; 95%CI, 0.68-1.65). There were no associations with any CVD in female patients (RR, 1.88; 95%CI, 0.43-8.24) and in those with preexisting CVD (RR, 1.31; 95%CI, 0.80-2.16). Heterogeneity between studies was high and significant except for the analysis on cerebrovascular diseases. This meta-analysis suggests no statistically significant association between ADHD medications and the risk of CVD across age groups, although a modest risk increase could not be ruled out, especially for the risk of cardiac arrest or tachyarrhythmias. Further investigation is warranted for the cardiovascular risk in female patients and patients with preexisting CVD as well as long-term risks associated with ADHD medication use.

Reference: Zhang L, Yao H, Li L, Du Rietz E, Andell P, Garcia-Argibay M, D'Onofrio BM, Cortese S, Larsson H, Chang Z. Risk of Cardiovascular Diseases Associated With Medications Used in Attention-Deficit/Hyperactivity Disorder: A Systematic Review and Meta-analysis. JAMA Netw Open. 2022 Nov 1;5(11):e2243597.

# iii) ADHD as a risk factor for TD2

We conducted a systematic review and a meta-analysis to quantitatively summarize evidence on the association between ADHD and T2D. Moreover, a register-based sibling study was conducted to simultaneously control for confounding factors. A systematic search identified four eligible observational studies (N = 5738,287). The meta-analysis showed that individuals with ADHD have a

more than doubled risk of T2D when considering adjusted estimates (OR=2.29 [1.48–3.55], *d*=0.46). Results from the register-based Swedish data showed a significant association between ADHD and T2D (HR=2.35 [2.14–2.58]), with substance use disorder, depression, and anxiety being the main drivers of the association, and cardiovascular and familiar risk playing a smaller role. While results from the meta-analysis provide evidence for an increased risk of T2D in individuals with ADHD, the register-based analyses show that the association between ADHD and T2D is largely explained by psychiatric comorbidities. Pending further evidence of causal association, our findings suggest that early identification and treatment of ADHD comorbidities might greatly reduce the risk of developing T2D in individuals with ADHD.

Reference: Garcia-Argibay M, Li L, Du Rietz E, Zhang L, Yao H, Jendle J, Ramos-Quiroga JA, Ribasés M, Chang Z, Brikell I, Cortese S, Larsson H. The association between type 2 diabetes and attentiondeficit/hyperactivity disorder: A systematic review, meta-analysis, and population-based sibling study. Neurosci Biobehav Rev. 2023 Apr;147:105076.

# 3. Conclusion

We have used state-of-the-art methods to synthesize and appraise the available evidence on ADHD and cardiometabolic disease with a focus on CVDs and T2D.

# APPENDIX: TIMESPAN publications for systematic review covering the available evidence on ADHD and cardiometabolic disease

- i) Li L, Yao H, Zhang L, Garcia-Argibay M, Du Rietz E, Brikell I, Solmi M, Cortese S, Ramos-Quiroga JA, Ribasés M, Chang Z, and Larsson H (In press). Attention-deficit/hyperactivity disorder is associated with increased risk of cardiovascular diseases: a systematic review and meta-analysis. JCPP Advances.
- ii) Zhang L, Yao H, Li L, Du Rietz E, Andell P, Garcia-Argibay M, D'Onofrio BM, Cortese S, Larsson H, Chang Z. Risk of Cardiovascular Diseases Associated With Medications Used in Attention-Deficit/Hyperactivity Disorder: A Systematic Review and Meta-analysis. JAMA Netw Open. 2022 Nov 1;5(11):e2243597.
- Garcia-Argibay M, Li L, Du Rietz E, Zhang L, Yao H, Jendle J, Ramos-Quiroga JA, Ribasés M, Chang Z, Brikell I, Cortese S, Larsson H. The association between type 2 diabetes and attention- deficit/hyperactivity disorder: A systematic review, meta-analysis, and population-based sibling study. Neurosci Biobehav Rev. 2023 Apr;147:105076.

# Attention-deficit/hyperactivity disorder is associated with increased risk of cardiovascular diseases: a systematic review and meta-analysis

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## 71 Abstract

Attention-deficit/hyperactivity disorder (ADHD) often co-occurs with other psychiatric and physical diseases. However, available evidence on associations between ADHD and cardiovascular diseases (CVDs) is mixed. To systematically review, quantitatively synthesize, and appraise available evidence on the link between ADHD with CVDs, we searched relevant articles in PubMed, Embase, PsycINFO, and Web of Science from inception to May 1, 2022. Study quality was assessed by using the Newcastle-Ottawa Scale (NOS), and random-effects model meta-analyses were performed. A total of 18,391,169 (ADHD: n=421,224) individuals from 11 studies were included in our systematic review and 8,196,648 (ADHD=332,619) individuals from five studies were included in the main meta-analysis of adjusted estimates. Pooled estimates showed that ADHD was significantly associated with an increased risk of CVDs in analyses based on adjusted effect size (OR=1.96; 95% CI =1.19-2.23, Q=140.74,  $P_0 < 0.001$ ,  $I^2 = 97.2\%$ ). When restricted among adults, the heterogeneity declined to null (OR=1.73; 95% CI =1.14-2.62, Q=6.28, Po=0.10, I<sup>2</sup>=6.28%), suggesting age might be the main source of heterogeneity. In subgroup analyses, we found increased risk of CVDs associated with ADHD across age groups, type of CVDs, and data sources. This systematic review and meta-analyses indicate that ADHD is associated with increased risk for CVDs, but further studies with various study designs are warranted to advance the understanding of the underlying mechanisms for the observed association between ADHD and CVDs. Additional research is also needed to resolve the role of ADHD medications which remains unclear due to the limited number of primary studies exploring this issue.

#### 106 Introduction

Cardiovascular diseases (CVDs) are the leading cause of morbidity, mortality, and 107 rising health care costs worldwide.(1, 2) CVDs caused an estimated 18.6 million deaths 108 in 2019 worldwide, corresponding to almost 400 million years of life lost and another 109 34.4 million years lived with disability.(3) It is therefore a public health priority to gain 110 111 further insight into the factors that contribute to CVDs in order to guide prevention and 112 clinical practice. Besides traditional CVD risk factors, such as overweight/obesity, diabetes, and smoking, concerns around the risk for CVDs in individuals with 113 psychiatric disorders are growing.(4-8) 114

Attention-deficit/hyperactivity disorder (ADHD) is one of the most prevalent 115 psychiatric disorders, affecting around 5% of children and 2.5% of adults.(9, 10) It is 116 characterized by developmentally inappropriate, pervasive and impairing inattention 117 118 and/or hyperactivity-impulsivity. In addition to the core clinical symptoms of ADHD, co-occurring poor mental and physical comorbidities are prevalent in individuals with 119 ADHD.(11) However, compared with the extensive research of psychiatric 120 comorbidities in ADHD,(12) physical comorbidities (13), such as CVDs, have received 121 122 less attention, particularly among adults. (14)

There are several reasons why individuals with ADHD could be at increased risk 123 for CVDs. First, previous research suggests that the link between ADHD and 124 125 cardiovascular diseases is biologically plausible via immune system abnormalities,(15, 16) neuromodulator dysregulation, (17, 18) and dysregulation of the hypothalamic-126 pituitary-adrenal (HPA) axis. (19, 20) Second, individuals with ADHD are at increased 127 risk for unhealthy lifestyle factors (e.g., smoking, obesity, poor physical activity), (21-128 23) which are all well-established risk factors for CVDs. (24-26) Third, previous studies 129 130 indicate that several psychiatric comorbidities of ADHD, for example, depression,(27, 28), substance use disorders, (29, 30) schizophrenia,(31) bipolar disorder,(32) and 131 anxiety disorder, (33) are associated with CVDs. Fourth, even though findings are 132 inconclusive, there has been a long-standing concern around a potential increased risk 133 of cardiovascular events due to the use of stimulant medications to treat ADHD. (34, 134 35) However, the available epidemiological findings are mixed, despite the potential 135 hypotheses for an increased risk of CVDs in ADHD. 136

The aim of the current study was to conduct a systematic review and meta-analysis was to provide a quantitative summary of observational studies on the associations between ADHD and CVDs. A second aim was to evaluate the impact of study design and confounder adjustment (e.g., sociodemographic factors, traditional risk factors for 141 CVDs, and use of medications) on the observed associations, in order to provide insights142 into underlying mechanisms.

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#### 144 Methods

The study was conducted and reported following the Preferred Reporting Items for
Systematic Reviews and Meta-Analyses (PRISMA) guidelines (36) and Conducting
Systematic Reviews and Meta-Analyses of Observational Studies of Etiology
(COSMOS-E). (37) This protocol was registered in the International Prospective
Register of Systematic Reviews (PROSPERO: CRD42021274367). (38)

150 Search Strategy

151 A systematic search for observational studies was conducted in PubMed, Embase, PsycINFO, and Web of Science databases, up to May 1, 2021, with no language and 152 article type restrictions. We used various combinations of the following keywords 153 "cardiovascular disease", "coronary heart disease", "heart disease", "sudden death", 154 "ischemic heart disease", "hypertension", "cerebrovascular disease", "stroke", 155 "transient ischemic attack", "attention-deficit hyperactivity disorder", "central nervous 156 system stimulants" and "observational study". The complete search strategy is 157 presented in Table S1. In addition, we performed manual searches through the reference 158 lists of original publications and reviewed articles to identify further pertinent studies. 159

#### 160 Inclusion criteria

We included all types of observational (cross-sectional or prospective) studies 161 providing data on the strength of the association between ADHD and CVDs in children, 162 adolescents, or adults. Eligible definitions of ADHD were as follows: 1) a categorical 163 diagnosis of ADHD according to the DSM-III, DSM-III-R, DSM-IV, DSM-IV-TR, 164 DSM-5, ICD-9, or ICD-10; 2) ADHD-medication prescriptions as a proxy for ADHD 165 diagnosis; 3) ADHD symptoms based on value above cut-off on a validated 166 self/parents/teacher-reported ADHD questionnaire; 4) ADHD diagnosed via a 167 structured psychiatric interview or positive answer by self/parents/teachers to the 168 question 'Did your doctor ever tell you that you/the child have ADHD?' Titles, abstracts, 169 170 and full text of included studies were screened independently by two authors. Discrepancies were resolved through discussion with a senior investigator. 171

The main outcome was the maximally adjusted odds ratio (OR), risk ratio (RR), and hazard ratio (HR), with their corresponding 95% confidence interval (CI) ratio expressing the association between ADHD and CVDs. The secondary outcome was the unadjusted estimate of the associations. When sufficient data (i.e., sample size,

prevalence of ADHD, CVDs) was available, crude ORs were manually calculated or 176 obtained by contacting the original authors, if not reported in the original paper. The 177 choice of primary and secondary outcomes was made because adjusted estimates are 178 more informative and potentially less prone to biased derived from the selection of the 179 participants.(39) 180

#### 181 **Data extraction**

The following information was extracted from each study for the qualitative and 182 quantitative synthesis: name of the first author, year of publication, sample size, data 183 source, study country, age and sex for participants, study design, years of original data 184 collection, the definition of ADHD and CVDs, effect size (ORs, RRs and HRs) and 185 186 information on confounding adjustment. We extracted both adjusted and unadjusted ratios if available, and we used the maximally adjusted ratios in the analyses. 187 Unadjusted effect sizes were calculated based on the information provided in the paper 188 when necessary. Two authors conducted the data extraction process separately, and any 189 disagreements were resolved by discussing with a third investigator. 190

The quality assessment of eligible studies was performed using the Newcastle-191 Ottawa Scale (NOS) by two independent authors, and the discrepancies were solved by 192 consensus. This 9-star scale consists of three parts: selection of participants and the 193 measurement of exposure (4 stars), comparability (2 stars), and assessment of outcomes 194 and adequate follow-up (3 stars). (40) A higher score on the NOS represents a higher-195 196 quality study, and generally, a score of 0 to 3, 4 to 6, or 7 to 9 is regarded as low-, moderate-, or high-quality, respectively. (40) 197

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#### **Statistical analysis**

We first described the characteristics of the included studies and various 199 200 confounding adjustment strategies to provide an overall picture of the current evidence. Meta-analyses were conducted using random-effects models in order to take into 201 account heterogeneity between studies, and the results were summarized in forest plots. 202 When there were several studies from the same population, only the one with the largest 203 sample size was included in the meta-analysis to avoid overrepresentation bias. (41) 204 205 ORs from logistic regression and HRs from Cox regression were combined because they closely approximate each other. (42) We then meta-analyzed adjusted and 206 unadjusted ORs across all studies. Additionally, several subgroups and sensitivity 207 analyses were conducted to investigate whether the main results were robust across (1) 208 age groups (adults and children separately) (2) specific types of CVD (hypertension 209 and other types of CVDs), and (3) different data sources (registers and research 210

samples). 211

Heterogeneity across studies was described by Cochran's Q test and the 212 inconsistency index  $(I^2)$ . If significant heterogeneity was detected by the Q test, we 213 considered I<sup>2</sup> values greater than 75% as high heterogeneity.(43) Restricted maximum 214 likelihood method was used to estimate between-study variability, with Hartung-215 216 Knapp-Sidik-Jonkman confidence interval for the summary effect. To evaluate each study's effect on the overall effect size, a leave-one-out analysis was also conducted. 217 The publication bias was first assessed through visual inspection of the funnel plot and 218 then tested quantitatively with Egger's test. All analyses were performed with Stata 16.0 219 (StataCorp L.P., College Station, TX). 220

221 Results

#### **Study Characteristics** 222

The study selection process is shown in Figure 1, and the list of excluded articles 223 (with reasons) after the full-text screen is presented in Table S2. Table 1 shows the 224 main characteristics of the eleven original studies included in the systematic review.(44-225 226 54) Among the 18,391,169 participants from six countries, a total of 421,214 individuals were with ADHD. The age of participants ranged from 5 to 81 years old, 227 and 9,889,671 (53.8%) participants were men. The publication years were between 228 2011 and 2022. The most common study design included in the meta-analysis were 229 cross-sectional (n=4), (44, 47, 49, 52), cohort (n=4) (46, 48, 50, 54), followed by case-230 231 control (n=3) (45, 51, 53) studies. These studies were conducted in the United States, Sweden, Germany, Turkey, Netherlands, and Canada. Of the 11 studies, five (44, 46, 232 47, 50, 54) obtained data from healthcare/insurance registries, five (45, 48, 49, 51, 53) 233 obtained data from research samples, and one (52) used data from a specific national 234 cohort. Therefore, register-based or electronic health care databases comprising large 235 numbers of participants were the most commonly used data source for the studied 236 associations. The definition of ADHD varied across studies, self/teacher/parent-237 reported ADHD symptoms (46, 48, 49, 52, 53) were the most widely used 238 measurements, followed by electronic health records based on codes in the ICD codes 239 240 or the DSM-IV (44, 45, 47, 50, 54). With regards to the CVD outcomes, various measurements were used, including CVD diagnosis based on ICD codes and several 241 specific recommendations/guidance for hypertension (e.g. Recommendations of the 242 American Heart Association). The most commonly studied CVD outcomes were 243 hypertension (44-47, 50, 51) and any type of CVDs (48, 49, 52-54). 244

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As shown in Table 2, seven (44, 46, 47, 49, 50, 52, 54) studies (64%) accounted

for potential confounders by adjusting for a number of measured covariates. However, 246 inadequate adjustment for confounding was commonly found among included studies. 247 Four studies (47, 49, 50) only adjusted for age and gender, while the other three studies 248 (46, 52, 54) also included sociodemographic characteristics (e.g., ethnicity, education), 249 psychiatric comorbidities, lifestyle factors (e.g., smoking, alcohol use, and physical 250 activity), and metabolic conditions (e.g., obesity, type 2 diabetes, and hyperlipidemia). 251 However, these lifestyle factors, psychiatric comorbidities, and metabolic conditions 252 are more likely mediators in the pathways linking ADHD to CVDs (Figure 2) rather 253 than confounders. The potential role of familial factors was only explored in one 254 Swedish study. (50) Of particular importance, only one study examined the contribution 255 256 of ADHD medication to the association between ADHD and CVDs in a sensitivity analysis, and suggested null effect of ADHD medication. (54) 257

The quality scores based on the NOS ranged from 2 to 9 stars (Median: 5), suggesting an overall moderate quality of the included studies. As shown in **Table S3**, the number of stars represented the score of each item. Generally, most studies used well-defined exposures and outcomes, but some included studies with one or no star in relation to 'Comparability' did not adjust for sex, age, and other sociodemographic characteristics.

#### 264 Meta-analysis

In the main analysis, we first explored the associations between ADHD and CVDs 265 among seven studies that provided adjusted effect sizes (44, 46-48, 50, 52, 54). The 266 three Swedish register-based studies (47, 50, 54) involved similar population, therefore 267 only the largest one was included in the main analysis. (47) A total of 8,196,648 268 (ADHD=332,619) individuals from five studies were included in the main meta-269 analysis of adjusted estimates. We found that ADHD was associated with a significantly 270 increased risk of CVDs (pooled odds ratio (OR) =1.96; 95% confidence interval (CI) 271 =1.19-2.23) (Figure 3). However, heterogeneity was high and significant (Q=140.74, 272  $P_0 < 0.001$ ,  $I^2 = 97.2\%$ ) and the between-study standard variance was not high (tau<sup>2</sup> is 273 0.09, 95%CI=0.03-0.44) but with wide confidence interval. The effect size was robust 274 275 in the leave-one-out sensitivity analysis (Figure S1), and the effect size was not driven by one single study. There was no evidence of publication bias for the primary outcomes 276 (Egger test: P=.81) (Figures S2 and S3). 277

When we repeated the analysis among six studies (44-46, 48, 51, 53) with unadjusted associations (five of six estimates were manually calculated), a positive association was found for CVDs in individuals with ADHD compared to controls (OR= 281 1.48, 95% CI =0.97-2.26; Q=188.57,  $P_Q < 0.001$ , I<sup>2=</sup>97.3%) but this was not statistically 282 significant (**Figure S4**).

Table 3 summarizes the results from subgroup and sensitivity analyses based on 283 adjusted estimates. First, in the analysis limited to adults, the association between 284 ADHD and CVDs was significant among six studies (pooled OR =1.73; 95% CI =1.14-285 286 2.62) without significant heterogeneity (Q=6.28,  $P_Q=0.10$ , I<sup>2</sup>=6.28%), while the only study in children reported a stronger association between ADHD and CVD 287 (hypertension as main outcome) than adults (OR =3.26; 95% CI =3.00-3.55; Qbetween-288  $_{\text{group}} = 86.74$ , P <0.0001). As age was a main factor affecting the heterogeneity of the 289 results, we restricted the following sensitivity analyses on type of CVDs and data 290 291 sources among adults. Second, among four studies with adjusted data on specific types of CVDs, we found that ADHD was associated with a higher risk of hypertension 292  $(OR=1.55; 95\% CI = 0.75-3.21, Q=5.89, P_0=0.06, I^2=66.0\%)$  and other CVDs 293  $(OR=1.73; 95\% CI = 1.19-2.68, Q=3.20, P_0=0.20, I^2=37.7\%)$ , even though the 294 association between ADHD and Hypertension was not statistically significant. 295

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#### 297 Discussion

The meta-analysis provided evidence that ADHD was significantly associated with 298 an increased risk of CVDs, after adjusting for potential confounders. However, due to 299 the limited number of available studies and lack of information in the original studies, 300 it is currently unclear whether these associations can be explained by other confounders. 301 Therefore, there is a critical need for future studies on this understudied topic to advance 302 the understanding of the underlying mechanism for the observed association between 303 ADHD and CVDs, especially the role of ADHD medications and other mediating 304 factors, as mental comorbid conditions (i.e. substance use disorders or depression). A 305 better understanding of major mediating factors may inform clinical guidelines about 306 how to intervene on CVDs in ADHD. 307

Compared with evidence from meta-analysis for well-established risk factors for 308 CVDs (i.e., smoking, obesity, physical activity/sedentary behavior, diabetes, 309 dyslipidemia and sleep disorders) among adults, the observed magnitude of association 310 between ADHD and CVDs was slightly smaller than the magnitude of association for 311 sedentary behavior, (26) diabetes, (55) and smoking, (24) but stronger than the 312 magnitude for dyslipidemia, (56) obesity (25) and sleep disorders (57) (Table 4). When 313 comparing with other psychiatric disorders, the observed strength of associations 314 between ADHD and CVDs is largely similar to estimates of associations of 315

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schizophrenia (58) and substance use disorders (59) with CVDs, but stronger than
stress-related disorder, (60) depression, (61) bipolar disorder, (58) and anxiety disorders
(Table 4). (62) Therefore, it is important to call for enhanced clinical awareness of
cardiovascular risk among adults with ADHD.

In the original studies, the type and number of possible confounding factors 320 321 adjusted for varied across studies, and few studies provided a clear rationale for their covariate selection, which is a common problem in contemporary observational 322 research on mental health. (63) Therefore, the significant adjusted association should 323 also be interpreted with caution. In addition, inadequate adjustment for confounding 324 was found in most studies (e.g., only four studies (44, 46, 52, 54) considered other 325 326 potential confounders except for age and sex), and only one of the included studies adjusted for ADHD medication, (54) which may contribute to increased risk of CVDs. 327 (64, 65) Only one of the previous studies used a sibling comparison design (50) and 328 suggested that the observed associations could also be partly explained by shared 329 familial factors. Therefore, it is currently unclear whether any of the observed 330 association reflect a causal effect or confounding. Clearly, more research using different 331 designs, such as within-family analysis or instrumental variable design, are needed to 332 systematically adjust for a broad set of possible confounders and further explore the 333 underlying mechanisms. Support for a potential causal association of ADHD with 334 coronary heart disease (66) and stroke (67) has been observed in a recent Mendelian 335 Randomization study. In addition, to identify critical intervention targets, more research 336 is also needed on potential mediating factors, such as ADHD medication use, other 337 psychiatric comorbidities and metabolic diseases. 338

Heterogeneity was high and significant for the main analyses, indicating that the pooled OR cannot appropriately summarize results from all the included individuals' studies, but it has limited effect on the conclusion of a positive association between ADHD and CVDs. (68, 69) When we restricted the analyses to different age group (children and adults), the associations remained stable, but the degree of heterogeneity substantially decreased to null, suggesting that age might be an important factor affecting the heterogeneity of the results.

Using subgroup meta-analyses, we clarified the differences in associations between ADHD and different types of CVD outcomes. The available studies indicate that individuals with ADHD had a 73% higher risk of a broad range of CVDs than those without ADHD. (54) suggested that the strength of the associations was more pronounced for cardiac arrest, hemorrhagic stroke, and peripheral vascular disease/arteriosclerosis. (54) However, research from independent samples is needed to replicate their findings. On the other hand, previous studies indicating that ADHDrelated traits such as impulsivity, hostility, and time urgency/impatience, are associated with increased risk for hypertension.(46, 70) Consistently, we also found a potential elevated risk of hypertension among adults with ADHD, but this was not statistically significant. Therefore, more studies are needed to further explore the associations between ADHD and specific types of CVDs.

358

### 359 Strengths and Limitations

This is the first systematic review and meta-analysis, with 421,224 individuals 360 361 having ADHD, to assess the relationship between ADHD and CVDs. Additionally, we also conducted sensitivity and subgroup analyses to further evaluate the findings from 362 the main analyses. However, our results should be considered in the context of some 363 limitations. First, we reported unadjusted ORs as secondary outcomes, but most (83.3%) 364 of the unadjusted ORs were from manually calculated effect sizes based on available 365 information in the original studies, as these studies did not generate unadjusted effect 366 size of the association between ADHD and CVDs. Second, we attempted to reduce 367 publication bias by including both published and unpublished studies, but bias cannot 368 be ruled out completely. Third, most of the included studies were conducted in Europe 369 and the U.S., which limits the generalizability of the findings to other populations 370 across the world. Therefore, more studies are needed to examine the associations 371 between ADHD and CVDs using samples from different settings and regions. Fourth, 372 the definition and measurements of ADHD and CVDs from original studies varied 373 substantially. Due to the limited number of included studies, it is not possible to assess 374 the associations among different definitions, which need to be evaluated in future 375 systematic reviews and meta-analyses on this topic. 376

#### **Future perspective**

Compared with other psychiatric disorders (e.g., schizophrenia, major depression), 378 the risk of CVDs in individuals with ADHD is largely understudied. A number of 379 380 important research questions, therefore, need to be addressed in future research. First, except for studies with a specific focus on hypertension, the most commonly used 381 measure of CVDs was a broad category, encompassing a wide range of circulatory 382 system diseases. Future studies are needed to explore the associations between ADHD 383 384 and specific types of CVDs, which is critical to enabling risk reductions via targeted intervention and preventative efforts. Second, future research should also consider 385

careful adjustments for a wide range of possible confounders of the observed 386 associations between ADHD and CVDs, including but not limited to early life risk 387 factors (e.g., preterm birth (71, 72) and birth weight, (72, 73)) and socioeconomic status, 388 (74, 75) as well as unmeasured familial factors. (50) To further identify critical 389 intervention targets, future research also needs to explore the role of potential mediating 390 391 factors, such as smoking, (76, 77) sleep problems, (78, 79) metabolic conditions (80) (e.g., obesity,(81) types 2 diabetes mellitus (T2DM),(82) dyslipidemia, (83) and in 392 particular obesity given that a recent Mendelian Randomization study suggested that 393 obesity may mediate the causal association between ADHD and coronary heart disease. 394 (66) Third, only one study has explored the role of psychiatric comorbidities and the 395 396 use of psychotropic medications in the studied associations. This is an important limitation given that ADHD is frequently comorbid with other psychiatric disorders 397 (e.g., mood disorder and substance use disorder), and these conditions and related 398 medications are in turn associated with increased risk of CVDs. (54, 84) Future studies 399 should examine to what extent the observed associations between ADHD and CVDs 400 could be explained by ADHD medications, other psychiatric disorders and related 401 medications. Fourth, research is also needed to test the potential sex- and age 402 differences in the associations of ADHD with CVDs, which is helpful for risk 403 stratification and individualized treatment recommendations in clinical practice. Taken 404 together, more studies on this topic are needed, especially using different study designs, 405 such as matched-cohort studies, (46, 50) genetically-informed studies (e.g., sibling 406 comparison studies and Mendelian Randomization studies) (66) and advanced 407 statistical methods (e.g. propensity score methods) to account for confounding. 408

#### 409 Conclusion

This systematic review and meta-analysis suggest a significant positive association between ADHD and CVDs. More efforts are needed to this substantially understudied research field. In particular, mediation effects by psychiatric comorbidities and related medications, as well as the causal mechanisms underlying the association, deserve further attention because of their important public health implications.

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#### 416 Conflict of Interest Disclosures

Dr Larsson has served as a speaker for Medice, Evolan Pharma and Shire/Takeda and
has received research grants from Shire/Takeda; all outside the submitted work. EDR
has served as a speaker for Shire Sweden AB outside the submitted work. Dr. Solmi

received honoraria/has been a consultant for Angelini, Lundbeck, Otsuka. S Cortese 420 declares honoraria and reimbursement for travel and accommodation expenses for 421 lectures from the following non- profit associations: Association for Child and 422 Adolescent Central Health (ACAMH), Canadian ADHD Alliance Resource 423 (CADDRA), British Association of Pharmacology (BAP), and from Healthcare 424 Convention for educational activity on ADHD. Dr. J. Antoni was on the speakers' 425 bureau and/or acted as consultant for Janssen-Cilag, Novartis, Shire, Takeda, Bial, 426 427 Shionogi, Sincrolab, Novartis, BMS, Medice, Rubió, Uriach, Technofarma and Raffo in the last 3 years. He also received travel awards (air tickets + hotel) for taking part in 428 psychiatric meetings from Janssen-Cilag, Rubió, Shire, Takeda, Shionogi, Bial and 429 Medice. The Department of Psychiatry chaired by him received unrestricted 430 educational and research support from the following companies in the last 3 years: 431 Janssen- Cilag, Shire, Oryzon, Roche, Psious, and Rubió. No other disclosures were 432 reported. 433

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671	FIGURES
672	Figure 1 PRISMA flow diagram for inclusion of the studies examining the association
673	between ADHD and cardiovascular diseases
674	PRISMA = Preferred Reporting Items for Systematic Review and Meta-Analyses
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677	Figure 2 Casual diagram representing the potential pathways of the association ADHD
678	and CVDs
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681	Figure 3 Forest plot of all studies describing associations between ADHD and CVDs
682	with adjusted estimates
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	Table 1 Overview of studies included in the systematic review												
Study	Country	Data sources and year of original	Study design	Sample	Ν	N (CVD)/N(NON-	Male	Mean age or	Exposure		Outcor	nes	NOS
Study	Country	data collection	Study design	size	(CVD)/N(ADHD)	ADHD)	(%)	age range	ADHD Definition	Age at assessment	Definition	Diseases	
Akmatov (2019) *	Germany	Insurance claims data, 2017	cross- sectional	2,586,620	4,837/258,662	16,519/2,327,958	75.6	5-14	ICD-10	5-14	ICD-10	Hypertension	7
Chen (2018) *	Sweden	National medical registers,1968- 2013	cross- sectional	5,551,807	2,072/61,129	247,870/5,490,678	50.8	40.55±13.49; 18-64	ICD-9 and 10	18-64	ICD-8, 9 and 10	Hypertension	6
Du Rietz (2021)	Sweden	National medical registers,1932- 2013	Cohort (follow for 47years)	4,789,799	NA/61,960	NA/4,727,839	51.0	18-81	ICD-9 and 10, ADHD medication prescription	N/A	ICD-8, 9 and 10	Hypertension, ischaemic heart disease, pulorary disease, atrial fibrillation, heart failure, stroke, peripheral vascular disease	8
Fuemmeler (2011) *	United States	National Longitudinal Study of Adolescent Health, 1995- 2009	Cohort (follow for 33years)	11,015	46/262	NA/10,753	51.0	28.80±0.12	Self-report	N/A	7th report of the Joint National Committee on Prevention Detection, Evaluation, and Treatment of High Blood Pressure: SBP $\ge$ 160 mm Hg and DBP $\ge$ 100	Hypertension	7
Grisaru (2018)	Canada	Research sample,2007-	case-control	3,804	3/55	53/3,749	51.1	6-19	Medical record	6-19	The fourth report on the diagnosis evaluation and	Hypertension	2

# Table 1 Overview of studies included in the systematic review

		2013									treatment of high blood	1	
											pressure in Children and	1	
											adolescents		
Li (2022)	Sweden	National medical registers,1941- 2013	Cohort (follow11.8 years)	5,389,519	663/37,027	62,089/5,352,492	48.9	38.44±12.32; 18-73	ICD-9 and 10	3-73	ICD-8, 9 and 10	CVD	9
Nilgün (2019)	Turkey	Research sample, 2012	case-control	177	10/77	9/100	24.3	8.95±2.68; 5-15	DSM-IV	N/A	Recommendations of the American Hear Association	e t Hypertension	3
Olazagasti (2013)	United States	Research sample	cohort	271	37/135	37/136	100.0	41.4±2.9	Teacher and parent rating	41.4±2.9	N/A	CVD	3
Semeijn (2013) *	Netherlands	Research sample,2008- 2009	cross- sectional	231	7/23	67/208	40.7	71.6±7.7	Semi-structured diagnostic interview	N/A	Self-report	CVD	5
Spencer (2014)	United States	Research sample	case-control	198	1/98	1/100	45.5	31±11	the adult ADHD self- reportscale (ASRS) v1.1 Symptom Checklist	N/A	Self-report	Heart attack	5
Xu (2021) *	United States	National Health Interview Survey, 2007 and 2012	cross- sectional	57,728	256/1,790	7,650/55,938	47.7	18 or older	Self-report	N/A	Self-report	All CVD Coronary heart disease Stroke	5

\*Included in the adjusted meta-analysis

		5											
		Covariates adjusted in each included study											
Study	Gender	Age/Year of birth	Region/race/ethnicity	Education achieved	Metabolic	Family income	Depression	Other	Smoking	Alcohol	Psychiatric	Family history	Physical activity
					conditions <sup>a</sup>			psychiatric			medications	of CVD	
								comorbidities					
								b					
Akmatov (2019)	×	×	×										
Chen (2018)	×	×											
Du Rietz (2021)	×	×											
Fuemmeler (2011)	×	×	×	×			×		×	×			×
Grisaru (2018)	N/A												
Li (2022)	×	×	×	×	×		×	×	×		×	×	
Nilgün (2019)	N/A												
Olazagasti (2013)	N/A												
Semeijn (2013)	×	×											
Spencer (2014)	N/A												
Xu (2021)	×	×	×	×		×			×	×			
``'													

Metabolic conditions include obesity, type 2 diabetes, dyslipidemia

B Other psychiatric comorbidities include anxiety disorder, autism spectrum disorder, bipolar disorder, conduct disorder, depressive disorder, eating disorder, intellectual disability, personality disorder, schizophrenia and substance use disorder

Study	Number of studies	Pooled ORs	Q	PQ	$I^2$
Adults	4	1.73(1.14-2.62)	6.28	0.1	6.28%
Children	1	3.26(3.00-3.55)	N/A	N/A	N/A
Adults:					
Hypertension	3	1.55 (0.75-3.21)	5.89	0.06	66.0%
Other CVDs	3	1.73(1.19-2.68)	3.2	0.201	37.7%

P<sub>Q</sub>: p-value associated to the Q statistic of heterogeneity

Table 4 The magnitude of associations between ADHD, well-established risk factors for CVDs, and psychiatric disorders with CVDs among adults

	OR/HR/RR (95% CI)	Type of evidence
ADHD	1.73(1.14-2.62)	Meta-analysis of observational studies
Well-established risk factors for CV	D	
Sedentary behavior (26)	2.47 (1.44-4.24)	Meta-analysis of observational studies
Diabetes (55)	2.29 (1.48-3.55)	Meta-analysis of observational studies
Smoking (24)	2.07 (1.82-2.36)	Meta-analysis of observational studies
Dyslipidemia (56)	1.43 (1.35-1.51)	Meta-analysis of observational studies
Obesity (25)	1.43 (1.33-1.54)	Meta-analysis of observational studies
Sleep disorders (57)	1.33 (1.13-1.57)	Meta-analysis of observational studies
Psychiatry comorbidities		
Schizophrenia (58)	1.91 (1.52–2.41)	Meta-analysis of observational studies

Substance use disorder (59)	1.70 ( 1.6 - 1.9)	Cohort study
Stress-related disorders (60)	1.64 (1.45-1.84)	Cohort study
Depression (61)	1.64 (0.84, 3.19)	Meta-analysis of observational studies
Bipolar disorder (58)	1.61 (1.34–1.94)	Meta-analysis of observational studies
Anxiety disorders (62)	1.52 (1.36 - 1.71)	Meta-analysis of observational studies



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### Abstract

**IMPORTANCE** Use of attention-deficit/hyperactivity disorder (ADHD) medications has increased substantially over the past decades, but there are concerns regarding their cardiovascular safety.

**OBJECTIVE** To provide an updated synthesis of evidence on whether ADHD medications are associated with the risk of a broad range of cardiovascular diseases (CVDs).

DATA SOURCES PubMed, Embase, PsycINFO, and Web of Science up to May 1, 2022.

STUDY SELECTION Observational studies investigating the association between ADHD medications (including stimulants and nonstimulants) and risk of CVD.

DATA EXTRACTION AND SYNTHESIS Independent reviewers extracted data and assessed study quality using the Good Research for Comparative Effectiveness (GRACE) checklist. Data were pooled using random-effects models. This study is reported according to the Meta-analyses of Observational Studies in Epidemiology guideline.

MAIN OUTCOMES AND MEASURES The outcome was any type of cardiovascular event, including hypertension, ischemic heart disease, cerebrovascular disease, heart failure, venous thromboembolism, tachyarrhythmias, and cardiac arrest.

RESULTS Nineteen studies (with 3 931532 participants including children, adolescents, and adults; 60.9% male), of which 14 were cohort studies, from 6 countries or regions were included in the meta-analysis. Median follow-up time ranged from 0.25 to 9.5 years (median, 1.5 years). Pooled adjusted relative risk (RR) did not show a statistically significant association between ADHD medication use and any CVD among children and adolescents (RR, 1.18; 95% CI, 0.91-1.53), young or middle-aged adults (RR, 1.04; 95% CI, 0.43-2.48), or older adults (RR, 1.59; 95% CI, 0.62-4.05). No significant associations for stimulants (RR, 1.24; 95% CI, 0.84-1.83) or nonstimulants (RR, 1.22; 95% CI, 0.25-5.97) were observed. For specific cardiovascular outcomes, no statistically significant association was found in relation to cardiac arrest or arrhythmias (RR, 1.60; 95% CI, 0.94-2.72), cerebrovascular diseases (RR, 0.91; 95% CI, 0.72-1.15), or myocardial infarction (RR, 1.06; 95% CI, 0.68-1.65). There was no associations with any CVD in female patients (RR, 1.88; 95% CI, 0.43-8.24) and in those with preexisting CVD (RR, 1.31; 95% CI, 0.80-2.16). Heterogeneity between studies was high and significant except for the analysis on cerebrovascular diseases.

**CONCLUSIONS AND RELEVANCE** This meta-analysis suggests no statistically significant association between ADHD medications and the risk of CVD across age groups, although a modest risk increase

(continued)

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#### **Key Points**

Question Are attention-deficit/ hyperactivity disorder (ADHD) medications associated with the risk of cardiovascular disease (CVD)?

Findings This systematic review and meta-analysis based on 19 observational studies with more than 3.9 million participants suggested that there was no statistically significant association between ADHD medications and the risk of cardiovascular events among children and adolescents, young and middleaged adults, or older adults.

Meaning Despite no statistically significant association between ADHD medications and CVD, more evidence is needed for the potential risk of cardiac arrest and tachyarrhythmias, the cardiovascular risk in female patients and in those with preexisting CVD, and long-term risk.

#### Invited Commentary

Supplemental content

Author affiliations and article information are listed at the end of this article

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#### Abstract (continued)

could not be ruled out, especially for the risk of cardiac arrest or tachyarrhythmias. Further investigation is warranted for the cardiovascular risk in female patients and patients with preexisting CVD as well as long-term risks associated with ADHD medication use.

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#### Introduction

Attention-deficit/hyperactivity disorder (ADHD), one of the most common neurodevelopmental disorders, is characterized by developmentally inappropriate inattention and/or hyperactivity-impulsivity symptoms starting in childhood.<sup>1</sup> The symptoms often persist into adulthood,<sup>2,3</sup> and even into older age for a substantial number of patients.<sup>4</sup> ADHD medications, including both stimulants and nonstimulants, are recommended for pharmacological treatment of ADHD, and the prevalence of ADHD medication use among both children and adults has increased substantially in many countries.<sup>5</sup>

While evidence from randomized clinical trials (RCTs) suggests ADHD medications are efficacious in reducing core ADHD symptoms,<sup>6</sup> there are concerns about their cardiovascular safety.<sup>7</sup> As ADHD medications are sympathomimetic agents that exert dopaminergic and noradrenergic effects, increasing heart rate and blood pressure is biologically plausible.<sup>8</sup> A previous Cochrane review of RCTs found that the stimulant methylphenidate was associated with increased pulse or heart rate.<sup>7</sup> However, as these RCTs could only evaluate short-term effects, it remains uncertain whether these changes led to a clinically significant risk of cardiovascular disease (CVD) over time. Longitudinal observation studies evaluating serious cardiovascular outcomes associated with ADHD medication use have emerged during the last decade, but with mixed findings.<sup>9-12</sup> A review paper incorporating five large population-based studies in the US reported no association between stimulants and serious cardiovascular events in children.<sup>13</sup> A meta-analysis of only 3 studies<sup>14</sup> found no increased risk of arrhythmic and ischemic cardiac events but a decreased risk of stroke. A more recent meta-analysis of 10 studies<sup>15</sup> showed a positive association between ADHD medications and risk of sudden death or arrhythmia but not for stroke, myocardial infarction, or all-cause mortality. However, it had several methodology limitations (eg, not preregistered, narrow outcome definition, and missing several important studies). Moreover, several new original studies have been published after these meta-analyses.<sup>16-20</sup> Thus, an updated synthesis is needed to address those limitations as well as to include a broader range of cardiovascular events (eg, hypertension, heart failure, and transient ischemic attack that have not been included in previous meta-analyses) and conduct sub-analyses by type of cardiovascular events and ADHD medications. In addition, observational studies that evaluate the benefits or risks of medical treatments are prone to bias (eg, immortal time bias, prevalent user bias, confounding by indication) if not conducted appropriately. It is therefore critical to make a rigorous quality assessment of the available studies and discuss common problems that future studies need to address. Understanding whether, and to what extent, ADHD medications are associated with CVD is highly relevant from both clinical and public health perspectives, as an increasing number of individuals are receiving ADHD medications globally. Findings of any significant association would prompt research on underlying causal mechanisms (eg, dopaminergic dysfunction and alterations in cytochrome P450 2D6 metabolism).<sup>8,21</sup>

The current study aims to provide a comprehensive and updated systematic review and metaanalysis to assess the associations between ADHD medications and risks of a broad range of cardiovascular events. In addition, we aim to examine whether there is any difference in the associations by types of ADHD medication, types of cardiovascular events, sex, age, and preexisting CVD conditions.

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#### **Methods**

This study was conducted and reported according to the Meta-analyses of Observational Studies in Epidemiology (MOOSE) checklist.<sup>22</sup> Our protocol is registered in the International Prospective Register of Systematic Reviews (CRD42021283702).<sup>23</sup>

#### Search Strategy and Selection of Studies

A systematic search for observational studies was conducted in MEDLINE via PubMed, Embase, PsycINFO, and Web of Science, up to May 1, 2022. We used various combinations of the following keywords: *cardiovascular disease, coronary heart disease, heart disease, sudden death, ischemic heart disease, hypertension, cerebrovascular disease, stroke, transient ischemic attack, attentiondeficit hyperactivity disorder, central nervous system stimulants, and observational study.* No restrictions to language were applied. The search strategy was designed with the assistance of a university librarian at Karolinska Institute (eTable 1 in the Supplement). In addition, we performed manual searches through the reference lists of relevant original publications and reviews to identify further pertinent studies.

We included all types of observational studies investigating associations between ADHD medication use and the risk of any CVD. We excluded reports, review articles, animal research, RCTs, and conference abstracts; studies without a comparator group; and studies with abuse or misuse of ADHD medication as the exposure. Titles, abstracts, and full text of included studies were screened independently by 2 investigators (L.Z. and H.Y.). Discrepancies were resolved through discussion with a senior investigator (L.L.).

#### **Data Extraction**

The following information was extracted from each study for the qualitative and quantitative synthesis: first author, year of publication, sample size, data source, study country, age and sex distribution, study design, year of original data collection, follow-up time, type of ADHD medication, measure of medication use, definition of CVD, relative risk, and covariate adjustment. Two investigators conducted the data extraction separately (L.Z. and H.Y.), and any disagreements were resolved through discussion with a senior investigator (L.L.).

Good Research for Comparative Effectiveness (GRACE) checklist version 2 was used for quality assessment.<sup>24</sup> Unlike the commonly used Newcastle-Ottawa Scale,<sup>25</sup> the GRACE checklist is tailored for evaluating the quality of observational studies that examine the outcomes of medical treatment. It evaluates the quality of observational research based on the use of concurrent comparators, equivalent measurement of outcomes in different groups, collection of data on confounders and effect modifiers, risk of immortal time bias, and reporting of sensitivity analysis.<sup>24</sup> Eleven items in the GRACE Checklist are grouped into 2 groups reflecting the quality of data and methods (eTable 2 in the Supplement). The quality assessment was completed by two investigators independently (L.Z. and H.Y.), and any discrepancies were solved by discussing with a senior investigator (L.L.).

#### **Statistical Analysis**

The characteristics of all included studies were described. Hazard ratios (HRs) from Cox regression, incidence rate ratios (IRRs) from Poisson regression, and odds ratios (ORs) from logistic regression were combined as approximations to relative risks (RRs), because under rare event assumption, different effect measures would yield mathematically similar estimates.<sup>26,27</sup> We used random-effects models to account for heterogeneity between studies. The significance of heterogeneity across studies was examined using Cochran *Q* test, while the percentage of variation attributed to true heterogeneity was estimated using the inconsistency index ( $l^2$ ).<sup>28</sup> The restricted maximum likelihood method was used to estimate between-study variability, with the Hartung-Knapp-Sidik-Jonkman confidence interval for the summary estimates.<sup>29,30</sup>

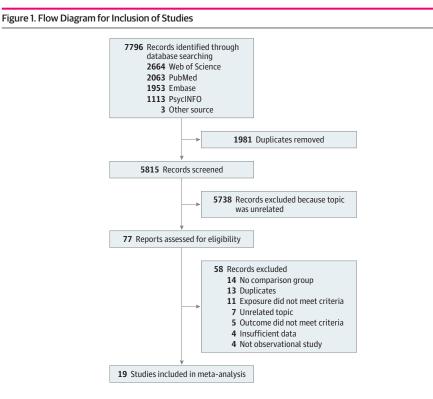
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We meta-analyzed adjusted RRs across all studies and by age groups (children and adolescents, young and middle-aged adults, and older adults). To evaluate each study's influence on the pooled estimates, the leave-one-out analysis was conducted. Publication bias was first assessed through visual inspection of the funnel plot and then tested quantitatively with Egger test. Subgroup analyses were conducted to investigate the associations of (1) stimulant and nonstimulant medications with any CVD, (2) ADHD medications with specific CVD (ie, cardiac arrest or tachyarrhythmias, cerebrovascular disease, myocardial infarction), (3) stimulant ADHD medications with specific CVD, (4) ADHD medications with any CVD in individuals with and without a history of CVD, and (5) ADHD medications with any CVD by sex. All analyses were performed with Stata version 16.0 (StataCorp). Statistical significance was set at *P* < .05, and all tests were 2-tailed.

#### Results

#### **Study Characteristics**

The process of study selection is shown in **Figure 1**. Detailed information on excluded articles with reasons is shown in eTable 3 in the **Supplement**. Overall, we included 19 studies published during 2007 to 2021, and their main characteristics are presented in **Table 1**.<sup>9-12,16-20,31-40</sup> A total of 3 931532 participants from 6 countries or regions (United States, South Korea, Canada, Denmark, Spain, and Hong Kong) were included. The study samples included children, adolescents, and adults, and 60.9% of participants were male. Average follow-up time ranged from 0.25 to 9.5 (median, 1.5) years. Most studies (14) were cohort studies, <sup>9-12,18,20,32-37,39,40</sup> followed by 3 nested case-control studies, <sup>16,19,31</sup> and 2 self-controlled case series. <sup>17,38</sup> The most common data source was insurance claims databases (15 studies). <sup>9,10,12,16,17,20,31-33,35-39</sup> More than half of the studies (10) used incident users, <sup>17,20,31-33,35-39</sup> while others included prevalent users. <sup>9-12,16,18,19,34,40</sup> Most studies used *International Classification of Diseases, Ninth Revision* or *International Statistical Classification of Diseases and Related Health Problems, Tenth Revision* to define CVD, except 1 study<sup>34</sup> that used self-reported CVD and another<sup>40</sup> without sufficient information on outcome measurement. Absolute



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Table 1. Charact	Table 1. Characteristics of Included Studies	uded Studies													
								Exposure				Outcomes			
Source (country)	Data source	Year of data collection	Study design	Median follow-up, y	Participants, No.	Male, %	Age, y	ADHD medication	New or prevalent user	Measure of use	Duration of use, y	Definition	Diseases	Adjustment for covariates	GRACE score
Cooper et al, <sup>9</sup> 2011 (United States)	Claims	1986-2005	Cohort	2.1	1 200 438	71	Mean, 11; range, 2-24	AMP, MPH, DMPH, ATX, pemoline	Mixed <sup>a</sup>	Current	NA	ICD-9, ICD-10	MI, CA, stroke	Demographics, comorbidities, health care utilization	11
Dalsgaard et al, <sup>11</sup> 2014 (Denmark)	Register	1990-2008	Cohort	9.5	8300	81	Mean, 11	AMP, MPH	Mixed <sup>a</sup>	Current	2.5	ICD-10	CVD	Demographics, comorbidities, perinatal characteristics	8
Guertin et al, <sup>31</sup> 2014 (Canada)	Claims	2001-2010	NCC	NA	38 495	70	Mean, 9	AMP, MPH, ATX	New	Current	NA	ICD-9, ICD-10	CVD	Demographics	Ø
Habel et al, <sup>10</sup> 2011 (United States)	Claims	1986-2005	Cohort	1.3	443 198	46	Range, 25-64	AMP, MPH, ATX, pemoline	Mixed <sup>a</sup>	Current	0.33	ICD-9, ICD-10	MI, CA, stroke	Demographics, CV risk score	11
Holick et al, <sup>32</sup> 2009 (United States)	Claims	2003-2006	Cohort	1.5	86 2 0 5	52	≥18	AMP, MPH, ATX	New	Current	NA	ICD-9	Cerebrovascular disease, TIA	Demographics, comorbidities, comedications	б
Houghton et al, <sup>16</sup> 2020 (United States)	Claims	2000-2016	NCC	NA	2046	68	Mean, 14; range, 3-18	AMP, MPH, ATX	Mixed <sup>a</sup>	Current	NA	ICD-9, ICD-10	MI, arrhythmia, stroke	Demographics, CV risk, comorbidities, comedications	6
Jeong et al, <sup>17</sup> 2021 (South Korea)	Claims	2002-2018	SCCS	NA	2104	51	Mean, 58; ≥6	HdM	New	Current	2.8	ICD-10	W	Age, comorbidities, comedications	10
Jeong et al, <sup>17</sup> 2021 (Taiwan)	Claims	2004-2015	SCCS	NA	484	64	Mean, 65; ≥6	НН	New	Current	2.3	ICD-9	MI	Age, comorbidities, comedications	10
Jeong et al, <sup>17</sup> 2021 (Hong Kong)	Clinical data	2001-2016	sccs	NA	30	50	Mean, 70; ≥48	НdМ	New	Current	2.3	ICD-9	M	Age, comorbidities, comedications	10
Latronica et al, <sup>18</sup> 2021 (United States)	EHR	2018-2020	Cohort	2	13233	37	Mean, 70; ≥65	AMP	Mixed <sup>a</sup>	Current	NA	ICD-10	HF, stroke, MI, AF, IHD, arrhythmia	Demographics, BMI, comorbidities	7
Olfson et al, <sup>33</sup> 2012 (United States)	Claims	1997-2007	Cohort	1.8	171 126	67	Range, 6-21	АМР, МРН	New	Current	NA	ICD-9	Angina pectoris, arrhythmia, TIA	Demographics, comorbidities, comedications	б
Peyre et al, <sup>34</sup> 2014 (United States)	National representative survey	2004-2005 e	Cohort	1	807	59	Mean, 40	Not specified	Mixed <sup>a</sup>	Ever	NA	Self-report	MI, angina pectoris, stroke	CV risk factors	ى ک
Saiz et al, <sup>19</sup> 2020 (Spain)	Primary care database	2002-2014	NCC	NA	2882	40	Mean, 14; range, 5-25	HdM	Mixed <sup>a</sup>	Current	0.65	ICPC, ICD-9	Valvular heart disease	Age, sex, smoking, comorbidities, comedications	6
Schelleman et al, <sup>35</sup> 2011 (United States)	Claims	1999-2006	Cohort	1.4	241 417	72	Range, 3-17	AMP, MPH, ATX	New	Current	0.37	ICD-9	Sudden death, ventricular arrhythmia	Data source	8
Schelleman et al, <sup>36</sup> 2012 (United States)	Claims	1999-2006	Cohort	1.2	219 954	45	≥18	HdM	New	Current	0.16	ICD-9	MI, stroke	Demographics, comorbidities, comedications	б
Schelleman et al, <sup>37</sup> 2013 (United States)	Claims	1999-2006	Cohort	1.2	192 905	46	≥18	AMP	New	Current	0.24	ICD-9	MI, stroke	Demographics, comorbidities, comedications	6
														) )	(continued)

Eutom         Eutom         Manue         Manue <th< th=""><th>Table 1. Charact</th><th>teristics of Incl</th><th>Table 1. Characteristics of Included Studies (continued)</th><th>continue</th><th>(þ</th><th></th><th></th><th></th><th></th><th></th><th></th><th></th><th></th><th></th><th></th><th></th></th<>	Table 1. Charact	teristics of Incl	Table 1. Characteristics of Included Studies (continued)	continue	(þ											
APH APH, APH, APH, APH, Cine									Exposure				Outcomes			
APH, APH, APH, APH, APH, APH, APH, APH,	Source (country)	Data source	Year of data collection	Study design				Age, y	ADHD medication	New or prevalent user	Measure of use	Duration of use, y	Definition	Diseases	Adjustment for covariates	GRACE score
APH APH, Ine APH, Ine Icine	Shin et al, <sup>38</sup> 2016 (South Korea)	Claims	2008-2011	sccs	NA	1224	78	Mean, 13; range, ≤17	МРН	New	Current	0.5	ICD-10	Arrythmias, HF, Hypertension, MI, stroke	Age, comorbidities, comedications	11
APH, APH, APH, icine	Tadrous et al, <sup>20</sup> 2021 (Canada)	Claims and health care database	2002-2016	Cohort		31310	49	Mean, 74; range, ≥66	АМР, МРН	New	Current	NA	ICD-9, ICD-10	MI, ventricular arrhythmia, stroke, TIA	Demographics, CVD history, physician visits, comorbidities, comedications	11
APH, APH, I	Winterstein et al, <sup>39</sup> 2007 (United States)	Claims	1994-2004	Cohort	2.3	55383	70	Range, 3-20	AMP, MPH, pemoline	New	Current	NA	ICD-9	MI, hypertensive diseases, aortic or thoracic aneurysm, arrhythmia, cardiac symptoms	Age, race, congenital anomalies, history of circulatory disease, comorbidities, comedications	б
APH,	Winterstein et al. <sup>12</sup> 2012 (United States)	Claims	1999-2006	Cohort	1.9	1219847	59	Range, 3-18	AMP, MPH	Mixed <sup>a</sup>	Current	NA	ICD-9	MI, CA, ventricular arrhythmia, stroke	Demographics, CV risk factors, comorbidities, and comedications	6
-	Zhang et al, <sup>40</sup> 2015 (United States)		1979-2014	Cohort	7.9	144	62	Mean, 11	AMP, MPH, ATX, guanfacine	Mixed <sup>a</sup>	Current	NA	NA	Syncope, aborted CA	Age, gender, QTc-duration, prior cardiac event, comedications	9
	Abbreviations: A atomoxetine; BN dexmethylpheni, Good Research fi <i>Revision; ICD-10</i> ,	DHD, attention II, body mass ir date: EHR, elec or Comparative International St	-deficit/hyperac ndex; CA, cardiac tronic health rec Effectiveness Ch tatistical Classific	:tivity disc c arrest; C :ord; HF, h hecklist; <i>I</i> ( <i>cation of L</i>	order; AF, atrial V, cardiovascul neart failure; EF CD-9, Internatic Viseases and Re	fibrillation; AMI ar; CVD, cardiov 4R, electronic h <i>i</i> <i>snal Classificatic</i> <i>slated Health Pr</i>	, amphe ascular ( salth rec in of Dise blems, i	tamines; / disease; DI ords; GRA( arses, Nini enth Revis	-	C, internation thylphenidat nemic attack. lixed indicate	al classific e; NA, not ; s mixture c	ation of prim available; NC of incident an	iary care; IHD, .C, nested casi .d prevalent u	ischemic heart disea: e-control: SCCS, self-c sers.	e; MI, myocardial infarction; ontrolled case series; TIA; tra	MPH, nsient

<sup>1</sup> JAMA Network Open. 2022;5(11):e2243597. doi:10.1001/jamanetworkopen.2022.43597

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risks of CVD in the included studies are shown in eTable 4 in the Supplement. All studies adjusted for measured covariates as an attempt to control for confounding, but the included covariates varied substantially across studies. The 2 self-controlled case series studies<sup>17,38</sup> further accounted for unmeasured confounders that are time invariant. The GRACE quality scores ranged from 5 to 11 (median, 9). Immortal time bias and lack of meaningful sensitivity analyses were the most common limitations in studies with lower scores (eTable 5 in the Supplement).

## **Meta-analysis Results**

We found that ADHD medication use was not statistically significantly associated with the risk of any CVD among children and adolescents (RR, 1.18; 95% CI, 0.91-1.53), young and middle-aged adults (RR, 1.04; 95% CI, 0.43-2.48), older adults (RR, 1.59; 95% CI, 0.62-4.05) (**Figure 2**), or overall (RR, 1.22; 95% CI, 0.88-1.68 (**Figure 3**). Analysis by effect measures also did not show significant estimates for HR (1.09; 95% CI, 0.84-1.42), OR (1.17; 95% CI, 0.51-2.66), or IRR (1.42; 95% CI, 0.43-4.68 (Figure 3). Heterogeneity between studies was high and significant (Cochran Q = 292.7; P < .001;  $I^2 = 93.2\%$ ). When restricting to specific effect measurements, heterogeneity was not significant for the analysis with HR as effect measures, yet it was still significant in other subgroups (Figure 3). As shown in the leave-one-out sensitivity analysis (eFigure 1 in the Supplement), the estimate was not driven by a single study. There was no evidence of publication bias, and a small study effect for the primary outcomes (eFigures 2 and 3 in the Supplement).

In subgroup analyses, we found no statistically significant associations of stimulant (RR, 1.24; 95% CI, 0.84-1.83) and nonstimulant medications (RR, 1.22; 95% CI, 0.25-5.97) with any CVD

Figure 2. Risk of Any Cardiovascular Event by Age Group Among Individuals Receiving Attention-Deficit/Hyperactivity Disorder Medication

		Lower risk Higher risk	14/ 1 L L L
Age group	RR (95% CI)	of CVD of CVD	Weight, %
Children and adolescents			
Shin et al, <sup>38</sup> 2016	0.98 (0.41-2.36)		4.30
Dalsgaard et al, <sup>11</sup> 2014	2.34 (1.15-4.75)		5.73
Guertin et al, <sup>31</sup> 2014	0.94 (0.82-1.07)		14.82
Houghton et al, <sup>16</sup> 2020	0.97 (0.68-1.39)		10.90
Jeong et al, 2021, <sup>17</sup> South Korea	1.43 (1.19-1.72)	<b>⊢</b>	14.08
Olfson et al, <sup>33</sup> 2012	0.69 (0.42-1.12)		8.58
Saiz et al, <sup>19</sup> 2020	0.28 (0.04-2.04)		1.09
Schelleman et al, <sup>35</sup> 2012	1.60 (0.19-13.60) 🔫		→ 0.94
Shin et al, <sup>38</sup> 2016	1.61 (1.48-1.74)		15.38
Winterstein et al, <sup>39</sup> 2007	1.20 (1.04-1.38)		14.71
Winterstein et al, <sup>12</sup> 2012	0.74 (0.38-1.46)		6.12
Zhang et al, <sup>40</sup> 2015	3.07 (1.09-8.64)		3.34
Subgroup: 1 <sup>2</sup> = 84.5%; P <.001	1.18 (0.91-1.53)	$\diamond$	100.00
Young and middle-aged adults			
Habel et al, <sup>10</sup> 2011	0.83 (0.72-0.96)		16.54
Jeong et al, 2021, <sup>17</sup> Hong Kong	19.76 (3.29-118.60)		→ 6.16
Jeong et al, 2021, <sup>17</sup> South Korea	1.38 (1.27-1.50)	-	16.67
Jeong et al, 2021, <sup>17</sup> Taiwan	0.39 (0.31-0.50)		16.23
Peyre et al, <sup>34</sup> 2014	1.57 (0.79-2.72)		13.90
Schelleman et al, <sup>36</sup> 2012	0.76 (0.50-1.18)		15.23
Schelleman et al, <sup>37</sup> 2013	0.78 (0.51-1.19)	<b>_</b>	15.27
Subgroup: <i>I</i> <sup>2</sup> = 95.5%; <i>P</i> <.001	1.04 (0.43-2.48)	:1	100.00
Older adults			
Jeong et al, 2021, <sup>17</sup> Hong Kong	5.60 (1.05-29.89)		● 9.15
Jeong et al, 2021, <sup>17</sup> South Korea	0.80 (0.74-0.87)	-	18.84
Jeong et al, 2021, <sup>17</sup> Taiwan	0.97 (0.86-1.10)	-	18.78
Latronica et al, <sup>18</sup> 2021	6.16 (4.22-8.99)		17.92
Schelleman et al, <sup>35</sup> 2012	1.16 (0.88-1.52)		18.37
Tadrous et al, <sup>20</sup> 2021	1.00 (0.60-1.80)		16.95
Subgroup: <i>I</i> <sup>2</sup> = 95.7%; <i>P</i> <.001	1.59 (0.62-4.05)		100.00
	0.2	1 RR (95% CI)	10

CVD indicates cardiovascular disease; RR, risk ratio.

(**Table 2**; eFigure 4 in the Supplement). When examining specific CVD outcomes (Table 2; eFigure 5 in the Supplement), no statistically significant associations were suggested for cardiac arrest or arrhythmias (RR, 1.60; 95% CI, 0.94-2.72), cerebrovascular diseases (RR, 0.91; 95% CI, 0.72-1.15), or myocardial infarction (RR, 1.06; 95% CI, 0.68-1.65). When examining stimulant medications, we found a similar pattern of results (eFigure 6 in the Supplement).

There was no association between ADHD medication use and any CVD for female (RR, 1.88; 95% CI, 0.43-8.24) and male (RR, 1.08; 95% CI, 0.32-3.67) patients. (Table 2; eFigure 7 in the

Figure 3. Risk of Any Cardiovascular Event by Measure of Association Among Individuals Receiving Attention-Deficit/Hyperactivity Disorder Medication

Measure	Effect measure (95% CI)	Lower risk of CVD	Higher risk of CVD	Weight, %
HR				
Cooper et al, <sup>9</sup> 2011	0.75 (0.31-1.85)			3.84
Dalsgaard et al, <sup>11</sup> 2014	2.34 (1.15-4.75)			4.44
Holick et al, <sup>32</sup> 2009	0.71 (0.34-1.47)			4.36
Schelleman et al, <sup>35</sup> 2011	1.60 (0.19-13.60) 🔫			→ 1.42
Schelleman et al, <sup>36</sup> 2012	1.01 (0.81-1.28)		-	5.81
Schelleman et al, <sup>37</sup> 2013	0.78 (0.51-1.19)		-	5.34
Tadrous et al, <sup>20</sup> 2021	1.00 (0.60-1.80)			4.96
Winterstein et al, <sup>38</sup> 2007	1.20 (1.04-1.38)		<b>.</b>	5.94
Zhang et al, <sup>39</sup> 2015	3.07 (1.09-8.64)			3.42
Subgroup: 1 <sup>2</sup> = 46.0%; P = .06	1.09 (0.84-1.42)	<	$\rightarrow$	39.54
OR				
Guertin et al, <sup>31</sup> 2014	0.94 (0.82-1.07)		-	5.95
Houghton et al, <sup>16</sup> 2020	0.97 (0.68-1.39)			5.52
Latronica et al, <sup>18</sup> 2021	6.16 (4.22-8.99)		—	- 5.47
Olfson et al, <sup>33</sup> 2012	0.69 (0.42-1.12)		-	5.15
Peyre et al, <sup>34</sup> 2014	1.57 (0.79-2.72)			4.74
Saiz et al, <sup>19</sup> 2020	0.28 (0.04-2.04) 🔫			1.61
Winterstein et al, <sup>12</sup> 2012	0.74 (0.38-1.46)			4.56
Subgroup: <i>I</i> <sup>2</sup> =93.6%; <i>P</i> <.001	1.17 (0.51-2.66)			32.99
IRR				
Habel et al, <sup>10</sup> 2011	0.83 (0.72-0.96)			5.94
Jeong et al, 2021, <sup>17</sup> Hong Kong	9.32 (3.44-25.28)			→ 3.53
Jeong et al, 2021, <sup>17</sup> South Korea	1.05 (1.00-1.11)		-	6.02
Jeong et al, 2021, <sup>17</sup> Taiwan	0.72 (0.65-0.80)	-8-		5.98
Shin et al, <sup>37</sup> 2016	1.61 (1.48-1.74)		-	6.00
Subgroup: 1 <sup>2</sup> =97.8%; P<.001	1.42 (0.43-4.68)			27.47
Overall: <i>I</i> <sup>2</sup> =93.2%; <i>P</i> <.001	1.22 (0.88-1.68)	<	$\diamond$	100.00
	0.2			10
		Effect n	neasure (95% CI)	

CVD indicates cardiovascular disease; HR, hazard ratio; IRR, incidence rate ratio; OR, odds ratio.

ysis Results			
Studies, No. (data sets, No.)	Pooled RRs	P <sup>a</sup>	I <sup>2</sup> , %
15 (17)	1.24 (0.84-1.83)	<.001	94.2
3 (3)	1.22 (0.25-5.97)	<.001	90.2
9 (9)	1.60 (0.94-2.72)	<.001	77.9
10 (10)	0.91 (0.72-1.15)	.14	34.0
8 (10)	1.06 (0.68-1.65)	<.001	86.2
3 (5)	1.08 (0.32-3.67)	<.001	96.1
3 (5)	1.88 (0.43-8.24)	<.001	85.6
10 (11)	0.99 (0.73-1.33)	<.001	98.7
7 (8)	1.31 (0.80-2.16)	<.001	99.0
	No. (data sets, No.) 15 (17) 3 (3) 9 (9) 10 (10) 8 (10) 3 (5) 3 (5) 10 (11)	Studies, No. (data sets, No.)         Pooled RRs           15 (17)         1.24 (0.84-1.83)           3 (3)         1.22 (0.25-5.97)           9 (9)         1.60 (0.94-2.72)           10 (10)         0.91 (0.72-1.15)           8 (10)         1.06 (0.68-1.65)           3 (5)         1.08 (0.32-3.67)           3 (5)         1.88 (0.43-8.24)           10 (11)         0.99 (0.73-1.33)	Studies, No. (data sets, No.)         Pooled RRs         P <sup>a</sup> 15 (17)         1.24 (0.84-1.83)         <.001

Abbreviations: ADHD, attention-deficit/hyperactivity disorder; CVD, cardiovascular disease; RR, risk ratio. <sup>a</sup> *P* value associated to the Cochran *Q* statistic of heterogeneity.

Supplement). We found no statistically significant associations in either individuals without a history of CVD (RR, 0.99; 95% CI, 0.73-1.33) or individuals with a history of CVD (RR, 1.31; 95% CI, 0.80-2.16) (Table 2). In particular, the only 2 studies<sup>11,40</sup> with long-term follow-up both showed elevated risk (RR, 2.01; 95% CI, 1.98-2.06 and RR, 3.07; 95% CI, 1.09-8.64) in those with a history of CVD (eFigure 8 in the Supplement).

# Discussion

## **Main Findings**

To our knowledge, this is the most comprehensive systematic review and meta-analysis of longitudinal observational studies on the association between ADHD medication use and the risk of CVD. By pooling results of 19 studies, we found no statistically significant association between ADHD medication use and CVD among children and adolescents, young and middle-aged adults, or older adults, although the pooled RR did not exclude a modest risk increase, especially for the risk of cardiac arrest or tachyarrhythmias. We did not detect any difference in the cardiovascular risk between stimulant and nonstimulant ADHD medication use. There was no association between ADHD medication and any CVD among female patients and those with preexisting CVD, although further study may be needed in these populations.

This updated meta-analysis enabled us to include 13 more studies than the previous meta-analyses.<sup>14,15</sup> Unlike the previous meta-analyses, we did not include all-cause mortality in our primary outcome (but sudden cardiac death), as previous studies have shown that most of the mortality in patients with ADHD was due to unnatural causes (eg, accidents and suicide).<sup>41,42</sup> We examined a broad range of cardiovascular outcomes including important cardiovascular outcomes (eg, hypertension and heart failure) in addition to those examined in previous meta-analyses (cardiac arrest, tachyarrhythmias, myocardial infarction, and stroke). We found no statistically significant association between ADHD medication use and CVD among children and adolescents, young and middle-aged adults, or older adults, although the confidence interval could not exclude an increased risk. It should be noted that as the absolute risk is relatively low, even a significant RR of 22% risk increase in general would possibly be offset by the benefits of medications, eg, alleviating ADHD symptoms and reducing risky behavior.<sup>6,43</sup> The trade-off between benefits and risks could be different in high-risk patients. Regarding specific cardiovascular outcomes, results from previous meta-analyses for specific cardiovascular outcomes (ie, cardiac arrest or tachyarrhythmias and stroke) are inconsistent.<sup>14,15</sup> We found that ADHD medication use seemed to be associated with an increased risk of cardiac arrest or tachyarrhythmias, but not with cerebrovascular disease and myocardial infarction.

We also reported several findings that were not explored in previous meta-analyses. In terms of types of ADHD medication, we found both stimulant and nonstimulant ADHD medications were not statistically significantly associated with any CVD, with similar pooled RRs. These would suggest similar null effects on CVD or similar degree of confounding in studies of both stimulants and nonstimulants. We were unable to compare stimulants vs nonstimulants for the risk of specific CVD due to the limited number of studies that examined nonstimulants. Of note, 1 previous open-label extension of an RCT study<sup>44</sup> compared the cardiovascular risks of a stimulant ADHD medication (dexmethylphenidate) vs a nonstimulant ADHD medication guanfacine. The study found that dexmethylphenidate was associated with increased systolic blood pressure, while guanfacine was associated with decreased heart rate, but both returned to baseline value during the 1-year open-label extension phase. It suggests that there might be differences in cardiovascular risks between stimulants and nonstimulants, but these differences may attenuate over time, thus not leading to a significant difference in clinically relevant outcomes. Nevertheless, head-to-head comparison studies based on observational data are warranted to compare stimulant vs nonstimulant ADHD medications regarding the risk of specific CVD.

We found that the risk of cardiovascular events associated with ADHD medications seemed to be higher among those with preexisting CVD compared with no prior CVD, although the findings did not reach the threshold for statistical significance. This coincides with raising concerns that individuals with congenital or acquired CVD are predisposed to additional risk.<sup>45</sup> Despite the lack of data supporting CVD history as a contraindication for ADHD medications, the FDA labeling includes a warning on the use of ADHD medications among individuals with structural cardiac abnormalities or other serious heart problems. Current treatment guidelines generally recommend carefully assessing patients with ADHD (eg, personal and family history of CVD, physical examination, electrocardiogram) and identifying individuals at risk before initiating ADHD medications.<sup>45</sup> Careful monitoring should also be performed after initiation.<sup>46,47</sup> Further studies focusing on the potential modifying risk of preexisting CVD, ideally separating risks for congenital or acquired CVD, are warranted. Clinical guidelines on prescribing ADHD medications among high-risk individuals should be updated once further evidence is available. We also found the point estimates for risk of CVD seemed to be higher among female compared with male patients, although only 3 studies have examined the sex-specific association along with high heterogeneity between studies. Previous research has shown that females with ADHD have somewhat different patterns of comorbidities<sup>48,49</sup> and response to stimulants<sup>50</sup> than males, and additional research is needed to examine this potential sex difference.

## **Methodology Consideration**

The analysis of observational data provides an emerging opportunity to generate evidence to inform clinical decisions, but there are important issues to consider to avoid biases.<sup>51</sup> One key issue is that treatment is not randomly assigned, which could result in confounded estimates. The included studies mainly reflected practice in clinical settings rather than controlled settings, so the prescription of ADHD medications is influenced by the clinician's perception of CVD risk. Most studies adjusted for a range of measured confounders, but the included confounders varied across studies. Many studies adjusted for demographic characteristics, and several adjusted for baseline comorbid conditions (eg, psychosis, obesity, and diabetes) and comedications (eg, antiepileptics, antidepressants, and asthma medications), yet few studies accounted for time-varying confounding factors. Moreover, several studies used general population control (rather than individuals with ADHD) as the comparison group, but only 1 study<sup>9</sup> adjusted for ADHD status. Not accounting for ADHD status would lead to bias, as recent research found that ADHD itself is a risk factor for CVD independent from comorbid psychiatric and somatic conditions.<sup>52</sup>

In addition, other fundamental flaws, such as selection bias and immortal time bias, need to be considered carefully when interpreting results from observational studies.<sup>51</sup> Nine of the 19 included studies<sup>9-12,16,18,19,34,40</sup> used prevalent users instead of incident users, and 7 studies<sup>16,18,19,34-37</sup> were at risk of immortal time bias. Misclassification and exclusion of the so-called immortality period would necessarily bias the results toward favoring the treatment.<sup>53</sup> Unlike lack of randomization, these flaws can be easily prevented by study design, eg, explicitly emulating a pragmatic target trial.<sup>54</sup> Moreover, most of the included studies (17 of 19) had an average follow-up time of up to 2 years. Only 2 studies had sufficient follow-up time to examine the long-term cardiovascular risk associated with ADHD medication, but these studies were only moderate in their quality (GRACE score, 6 and 8 of 11). Thus, further studies with rigorous methods are needed to evaluate the long-term risk of CVD associated with ADHD medication use.

## **Clinical Implications**

Overall, our meta-analysis provides reassuring data on the putative cardiovascular risk with ADHD medications, but the possible associations with cardiac arrest or tachyarrhythmias, among female patients, and among those with preexisting CVD warrants further investigation. Importantly, our findings are presented at the population level; in clinical practice, specific individuals with ADHD might be particularly prone to negative cardiovascular outcomes; therefore, clinicians should discuss

with their patients and families the possible cardiovascular risk of ADHD medication in light of the latest evidence, and they should rigorously follow clinical guidelines that suggest monitoring of blood pressure and heart rate at baseline and each medication review.

## Limitations

There are several limitations to consider when interpreting the results. First, heterogeneity was high and significant for most analyses. Although this heterogeneity does not invalidate our results, it indicates that the pooled RR cannot appropriately summarize results from all individual studies and should therefore be interpreted with caution.<sup>55</sup> When restricting to specific cardiovascular outcomes, heterogeneity was not significant for the analysis on CVDs, yet it was still significant in the subgroup analyses by sex and preexisting CVD. Second, due to a lack of data, we were unable to compare the associations with specific ADHD medications. Third, as few studies have information on dosage and duration of medication use, investigation of the dose-response association was not possible. Fourth, although the GRACE checklist is validated for evaluating the quality of observational studies of medical treatment, a total score approach for risk of bias assessment needed to be validated. Additionally, most of the included studies were conducted in the United States and Europe, which means the results may not generalize to other settings.

# Conclusions

The results of this meta-analysis suggested no statistically significant association between ADHD medication use and the risk of any cardiovascular events across age groups, although a modest risk increase could not be excluded, especially for the risk of cardiac arrest or tachyarrhythmias. Our study also warrants future studies with rigorous study designs to investigate the risk of cardiovascular events among female patients and among those with preexisting CVD, as well as the long-term risk of ADHD medication use.

#### **ARTICLE INFORMATION**

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**Author Contributions:** Ms Zhang and Dr Chang had full access to all of the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis. Mss Zhang, Yao, and Dr Li contributed equally.

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# Neuroscience and Biobehavioral Reviews

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# The association between type 2 diabetes and attention- deficit/ hyperactivity disorder: A systematic review, meta-analysis, and population-based sibling study

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ARTICLE INFO	A B S T R A C T
<i>Keywords:</i> Type 2 diabetes Attention deficit hyperactivity disorder Meta-analysis Cardiovascular risk factors	We conducted a systematic review and a meta-analysis to quantitatively summarize evidence on the association between attention-deficit/hyperactivity disorder (ADHD) and type 2 diabetes (T2D). Moreover, a register-based sibling study was conducted to simultaneously control for confounding factors. A systematic search identified four eligible observational studies (N = 5738,287). The meta-analysis showed that individuals with ADHD have a more than doubled risk of T2D when considering adjusted estimates (OR=2.29 [1.48–3.55], <i>d</i> =0.46). Results from the register-based Swedish data showed a significant association between ADHD and T2D (HR=2.35 [2.14–2.58]), with substance use disorder, depression, and anxiety being the main drivers of the association, and cardiovascular and familiar risk playing a smaller role. While results from the meta-analysis provide evidence for an increased risk of T2D in individuals with ADHD, the register-based analyses show that the association between ADHD and T2D is largely explained by psychiatric comorbidities. Pending further evidence of causal association, our findings suggest that early identification and treatment of ADHD comorbidities might greatly reduce the risk of developing T2D in individuals with ADHD.

#### 1. Introduction

Attention-Deficit/Hyperactivity Disorder (ADHD) is a prevalent neuropsychiatric condition characterized by age-inappropriate and impairing inattention and/or hyperactivity/impulsivity. Although once conceived as a childhood limited disorder, ADHD has been estimated to affect 5–10 % of school-age children worldwide and 2–5 % of adults (Polanczyk et al., 2014), and its impairing symptoms persist into adulthood in up to 65 % of those diagnosed with ADHD in childhood (Faraone et al., 2006). Previous research has provided evidence of significant psychiatric comorbidity in ADHD (Faraone et al., 2021), which has informed clinical guidelines to support the diagnosis and treatment

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of ADHD patients with co-occurring psychiatric disorders. Results from genome-wide association studies (Demontis et al., 2019), cohort studies (Galéra et al., 2022), and an umbrella review of meta-analyses of observational studies (Arrondo et al., 2022) suggests that ADHD is also associated with physical conditions. More specifically, there is meta-analytic evidence of a significant association between ADHD and obesity (Cortese et al., 2016), and ADHD and asthma (Cortese et al., 2018; Sun et al., 2021). However, evidence on the association between ADHD and a physical condition associated with obesity, namely type 2 diabetes mellitus (T2D), is sparse and has not been meta-analysed yet.

T2D is a metabolic disorder characterized by chronic hyperglycemia and insulin resistance with a worldwide rising prevalence that more than doubled over the past decades translating into a global economic burden and a serious public concern (Zhang and Gregg, 2017). T2D used to be considered as solely occurring in adults, while research has now demonstrated an increasing incidence in young adults, adolescents, and children (Pinhas-Hamiel and Zeitler, 2005). T2D affects individuals' quality of life and is associated with medical comorbidities and increased mortality (Pantalone et al., 2015). Risk factors for T2D include modifiable factors such as overweight and obesity, sedentary behaviors, poor dietary habits, smoking, hypertension, sleeping disorders, depression, antipsychotics use and non-modifiable factors such as age, family history of T2D, and history of gestational diabetes (Chen et al., 2012; Grajales et al., 2019).

Therefore, both ADHD and T2D share several risk factors and comorbidities that require careful consideration to advance the understanding of why ADHD and T2D co-occurs. Whilst the association of ADHD with T2D-related cardiometabolic comorbidities such as obesity, hypertension, and maternal diabetes is established (Cortese et al., 2016; Fuemmeler et al., 2011; Garcia-Argibay et al., 2022; Zhao et al., 2019), less is known on the association between ADHD and T2D. Although an association between ADHD and T2D has been documented in individual studies (Chen et al., 2013; Chen et al., 2018a, 2018b; Du Rietz et al., 2021; Xu et al., 2021), the magnitude of the association is not consistent across studies. Furthermore, no study has controlled simultaneously for a number of potential mediating factors including psychiatric comorbidities (e.g., anxiety disorders, depression, schizophrenia, substance use disorder) and unmeasured familial confounding (i.e., genetic and environmental risk factors shared by family members).

Therefore, the aim of this study was to 1) critically review all observational studies on the association between ADHD and T2D, 2) meta-analyze the available studies to establish the association between ADHD and T2D, and 3) address limitations of the current studies such as poor confounder/mediator control by conducting a large nationwide sibling study using national register data from Sweden that allowed us to simultaneously control for a large number of confounding factors. We hypothesize that psychiatric comorbidity in ADHD may explain part of the ADHD-T2D association. In particular, we expected that internalizing disorders such as anxiety and depression might increase the risk of developing T2D among people with ADHD.

#### 2. Methods

#### 2.1. Systematic review and meta-analysis

The study protocol was preregistered on the International Prospective Register of Systematic Reviews (PROSPERO; Number: CRD42022322364). We followed guidelines from the Meta-Analysis of Observational Studies in Epidemiology group (MOOSE; Stroup, 2000) and Preferred Reporting Items for Systematic reviews and Meta-Analysis (PRISMA) (Page et al., 2021).

## 2.1.1. Search strategy and selection of studies

A systematic literature search was performed in EMBASE, MEDLINE via PubMed and Web of Science, and databases were searched from their inception date until August 2nd, 2021. The search terms and syntax are

provided in the supplement. From the search results, we selected articles using the following criteria: 1) Studies assessing the relationship between ADHD and T2D in individuals of any age and sex; 2) peer reviewed papers; 3) observational studies (case-control or cohort studies), 4) diagnosis of ADHD defined as a) presence of a register-based diagnosis according to DSM (III, III-R, IV, IV-TR or 5) or hyperkinetic disorder according to ICD-9 or ICD-10 codes 314/F90; or ADHD medication prescriptions as a proxy to diagnosis or b) sum score of ADHD symptoms above an established cut-off based on validated rating scales such as Child Behavior Checklist (CBCL) and Strengths and Difficulties Questionnaire (SDQ), assessed by parents, teachers, or self-ratings, or c) a positive answer from the individual to a question similar to "have you ever been diagnosed with ADHD?" or by parents to the question: "Has the child ever been told having ADHD by a doctor?". When there were multiple studies from the same population, only the one with the largest sample size was included in the meta-analysis to avoid overrepresentation bias. Two authors (MG and LL) independently screened the articles for relevance and eligibility, and, in case of disagreement, discrepancies between authors were adjudicated by a third, senior reviewer (HL). Fig. 1 displays the PRISMA flowchart of the search, with respective reasons for study exclusion.

#### 2.1.2. Data extraction

Data extracted included: surname of the first author, year of publication, country, source of the data, age range, total sample size, number of cases and controls with ADHD, number of cases and controls with T2D, mean age of the sample, percentage of men, covariates adjusted for, and maximally adjusted odds ratio (OR), risk ratio (RR), or hazard ratio (HR).

#### 2.1.3. Assessment of study quality

Before performing the meta-analysis, a quality assessment was performed of all eligible studies. Study quality was evaluated using the Newcastle–Ottawa Scale (NOS; Stang, 2010), and study quality was scored from zero to nine on the basis of study group selection, comparability, and outcome and follow-up. NOS scores lower than 7 deemed low quality. Disagreements were resolved by consensus.

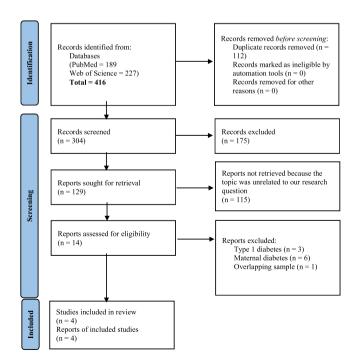


Fig. 1. PRISMA 2020 flowchart of the systematic review.

#### 2.1.4. Statistical analysis

Given the expected variability across studies, a random-effects model with a restricted maximum likelihood (REML) estimator and the Knapp-Hartung (IntHout et al., 2014; Knapp and Hartung, 2003) small-sample adjustment was fitted to calculate the pooled OR with 95 % confidence interval (CI). The RR and HR were considered as OR for the pooled analysis given the low prevalence of ADHD in adults and older adults (Zhang and Yu, 1998). Heterogeneity between studies was assessed using the Cochrane Q-test and the  $I^2$  statistic, with significant heterogeneity indicated when  $p_q < 0.1$  or  $\geq 50$  %, respectively (Higgins, 2003). Profiled restricted log-likelihood plots with respect to  $\tau^2$  were examined to ensure that a global maximum was found. In order to detect potential outliers/influential studies, studentized residuals and Cook's distances were used. Studies were considered outliers when studentized residuals were larger than  $100 \times (1 - 0.05/(2 \times k))th$  percentile of a standard normal distribution, and influential studies when a Cook's distance was larger than the median $+ 6 \times IQR$ . Methods to test for publication bias such as a funnel plot, Egger regression asymmetry test, or the adjusted rank correlation were not carried out given the small number of studies. We used a method that is not based on funnel plot asymmetry, namely a one-parameter selection model (Vevea and Woods, 2005) using a half-normal selection function. This method attempts to detect publication bias by modelling the underlying selection process by which the included studies in a meta-analysis might have been influenced and corrects the estimates (Sterne et al., 2001). The trim-and-fill method was used to assess the stability of the pooled results. Leave-one-out sensitivity analysis was performed in order to identify individual studies with a substantial influence on the between-study heterogeneity or overall risk estimate. In order to detect possible biases, effect sizes were regressed on NOS score, total sample size, and year of publication. A Bayesian linear regression with the inverse of the variance as weights was fitted using Jeffreys-Zellner-Siow (JZS) priors (Wetzels and Wagenmakers, 2012). The null model and the models including NOS score, total sample size, and year of publication were compared using Bayes Factor (BF) (Schönbrodt and Wagenmakers, 2018). Lastly, to ensure adequate statistical power (>0.80) for the pooled effect size and test of homogeneity, post hoc power analyses were carried out. All data analyses were performed in R 4.1.0 (R Development Core Team, 2020) with the r-package metafor version 3.1.46 (Viechtbauer, 2010).

#### 2.2. Nationwide population-based cohort study

The study population was based on the linkage of several populationbased national registers in Sweden linked by unique personal numbers (Ludvigsson et al., 2009), namely The Total Population Register (TPR), Cause of Death Register, Prescribed Drug Register (PDR), National Patient Register (NPR), the Multi-generation Register (MGR), and Longitudinal Integration Database for Health Insurance and Labor Market Studies (LISA).

The cohort included all individuals born between 1941 and 1983 who were alive and living in Sweden in 2001 with information on their biological parents (N = 4,257,955). Individuals with a previous history of T2D before the start of follow-up were excluded from the analyses (n = 40,795). We identified full siblings through the MGR (1,306,841 clusters with at least 2 full siblings). The study had ethical approval from the Regional Ethical Review Board in Stockholm, Sweden (Dnr 2013/ 862–31/5). The requirement for informed consent was waived because the study was register-based and data on the included individuals were deidentified. The investigation conforms to the 1964 Helsinki declaration and its later amendments or comparable ethical standards.

## 2.2.1. Exposure

Individuals with ADHD were identified as those who had either an ADHD diagnosis (inpatient or outpatient specialist care services) after age 3 years in the NPR using the International Classification of Diseases (ICD) version 9 (1987–1996; ICD-code 314) or ICD version 10 (1997present; ICD-code F90) or a prescription of any medication approved for the treatment of ADHD in Sweden during the follow-up period (methylphenidate: N06BA04, amphetamine: N06BA01, dexamphetamine: N06BA02, atomoxetine: N06BA09, lisdexamfetamine: N06BA12) from the PDR. Guanfacine (C02AC02) was not included as, in Sweden, it was not approved for the treatment of ADHD until 2016 (Huss et al., 2016).

#### 2.2.2. Outcome

T2D was defined as the presence of a registered diagnosis in the NPR of ICD version 10, ICD-code E11, ICD version 8 and 9, ICD-code 250.

#### 2.2.3. Covariates

Demographics such as year of birth, sex, and highest achieved education (elementary, high school, and postgraduate) were collected from the TPR and LISA. Diagnoses of psychiatric conditions including anxiety, depression, schizophrenia, bipolar disorder, and substance use disorder, as well as cardiovascular risk factors including hyperlipidemia, obesity, and sleep disorders were identified from the NPR using ICD codes (inpatient or outpatient specialist care services). Individuals prescribed antipsychotic drugs were identified as a registered prescription of any medication with Anatomical Therapeutic Chemical (ATC) codes N05A, excluding lithium (N05AN01). Antipsychotics were included given that they may be used in individuals with ADHD, especially for aggressiveness (Zhang et al., 2021), and its relationship with T2D (Galling et al., 2016). Each disorder/medication was defined as 0 or 1 based on whether the person has ever been diagnosed/prescribed with each disorder or drug before or during that person-time interval. To facilitate open science and transparent reporting (Larsson, 2022) all ICD and ATC codes used to define each covariate are presented in Supplementary Table S1. Missing data on education was dealt with by creating a new factor level for those with missing values to avoid listwise deletion from our models.

#### 2.2.4. Statistical analysis

All individuals were followed up from January 1st, 2001 —when outpatient data was introduced— until death, emigration, date of the first T2D diagnosis, or December 31st, 2013, whichever occurred first. Individuals with T2D before the start of follow-up were excluded. A time-dependent Cox regression model with age as the underlying time scale was fitted to assess the relationship between ADHD and T2D. Each psychiatric disorder (including ADHD) and physical condition was allowed to vary over time from unexposed to exposed (0/1), i.e., the hazard at time *t* depends only on the value of each covariate at that given timepoint.

The concordance index (C-index) and Bayesian Information Criterion (BIC) were used to measure the goodness-of-fit for each model and to compare fit to the unadjusted models. First, we performed a series of sequential adjustments: 1) a crude model (henceforth referred to as baseline) adjusted for birth year and sex, 2) we further adjusted for psychiatric conditions (i.e., anxiety, depression, schizophrenia, bipolar disorders, substance use disorder) and cardiometabolic risk factors (i.e., hyperlipidemia, obesity, sleep disorders) that could mediate this relationship, and 3) we adjusted for all aforementioned variables together with use of antipsychotic medications. Second, we stratified by sex in order to assess potential sex differences. Third, in separate models, we adjusted the baseline model for 1) education, 2) psychiatric comorbidities, 3) psychiatric comorbidities plus antipsychotic use, and 4) cardiometabolic risk factors and compared estimates to the baseline model to determine which mediator attenuates the association the most. Finally, to explore the extent to which this relationship is influenced by unmeasured familial confounders shared within sibling pairs, a Cox model adjusted for birth year and sex with a separated stratum for each cluster of full siblings was fitted. This method adjusts for shared familial confounders including genetic factors and shared unmeasured confounders (Allison, 2009). Benjamini–Hochberg correction was

performed to control false discovery rate due to multiple resting. Data management and statistical analyses were performed using SAS software version 9.4 (SAS Institute, Cary, NC) and R version 4.1.0 (R Development Core Team, 2020).

#### 2.2.5. Sensitivity analyses

A series of sensitivity analyses were carried out to assess the robustness of our results. First, analyses were rerun assuming that individuals who received an ADHD diagnosis at any point during follow up had ADHD since the start of follow up in 2001, in order to test the robustness of the results. Second, to explore the possibility of different magnitude of associations at different ages, we reran all baseline analyses stratified by age at the start of follow up divided into four categories (less than 31, 31–40, 41–50, and 51–72 years). Adjusted analyses were not performed when stratifying by age categories due to low count of individuals with both ADHD and T2D.

## 3. Results

## 3.1. Systematic review and meta-analysis

#### 3.1.1. Study characteristics

We identified four individual studies with non-overlapping samples that met our inclusion criteria for the meta-analysis. The total number of individuals was 5,738,287 (103,022 individuals with ADHD and 5,635,265 without ADHD). The mean age of the overall population was 26.88 years (SD = 18.82; range 5–79) and 64.53% were men. Among the four studies, three used register-based data (Chen et al., 2013; Chen et al., 2018a), and one used questionnaire data. For the diagnosis of ADHD and T2D, one study relied on retrospective recollection from a clinician (Xu et al., 2012), two studies used version ICD-9 codes (Chen et al., 2013; Chen et al., 2013; Chen et al., 2018a), and one used version ICD-9 and version ICD-10 codes (Chen et al., 2018b). A summary of the selected studies is shown in Table 1. The NOS rated all four studies as high quality, with an average NOS score of 8.75/9 (see Supplementary Table S2).

One register-based study from Taiwan (Chen et al., 2013), including 4,302 newly diagnosed ADHD patients and 21,510 randomly selected controls, displayed a significant association between ADHD and T2D (OR = 2.83) after adjustments. Individuals with ADHD had a higher prevalence than controls (0.8 vs 0.3%). Similarly, using the Taiwan National Health Insurance Research Database in a matched-control cohort design, 107,847 adolescents and young adults (35,949 individuals with ADHD and 71,898 age- and sex-matched controls), with mean age of 12.89 years and predominantly men (78.8%), were followed for up to 9 years (Chen et al., 2018a). The results showed an increased risk of developing T2D later in life in both adolescents and young adults (HR = 2.83 and HR = 3.28, respectively) after adjustments for demographic characteristics, ADHD medications, atypical antipsychotics, and comorbidities. With the linkage of multiple Swedish national registers, Chen et al. (2018b) explored the association between ADHD and T2D in a total sample of 5,551,807 adults (50.8% males) with a mean age of 40.55 years. In those with ADHD, the prevalence ratio (PR) was more than twice as large compared to those without ADHD (PR = 2.41) after adjustments for sex and age, with a higher prevalence in males (4.32% vs 3.58%). Lastly, Xu et al. (2021) using data from the National Health Interview Survey (NHIS; 2007 and 2012 cycles) in a study of 52,821 adults (48.6 % men) with a mean age of 45.5 years (range 20-79). They found an increased likelihood of having T2D among those with ADHD compared with those without ADHD after adjustments (OR = 1.54).

## 3.1.2. Meta-analysis

A total of k = 4 studies were included in the analysis. The profiled restricted log-likelihood plot displayed a unimodal distribution and that a global maximum was reached. The estimated average adjusted OR based on the random-effects model was  $\hat{\mu} = 2.29, 95 \%$  CI [1.48–3.55],

<b>Table 1</b> Study cha	uracteristi	Table 1Study characteristics of the included studies.	cluded stu	dies.												
Study	Year	N <sub>total</sub>	$\mathbf{N}_{\mathbf{adhd}}$	N <sub>adhd</sub> N <sub>noadhd</sub>	Mean age	% Men	Adjustments	Data type	Country	Age range	Diagnoses	IR ADHD	IR no ADHD	Raw IRR	aOR	SON
Chen HJ. et al.	2013	25,812	4302	21,510	8.6	80	Age, sex, index year, geographic location, and obesity	Register- based	WT	5-15	ICD-9	8.37 (5.87–11.57)	2.98 (2.29–3.80)	2.81 (1.87–4.23)	2.75	6
Chen MH. et al.	2018	107,847		35,949 71,898	12.9	78.8	Age, sex, level urbanization, income, ADHD medication, antipsychotic medication, hypertension, dyslipidemia, and obesity	Register- based	WT	1029	ICD9	4.40 (3.74–5.13)	1.10 (0.87–1.37)	4.01 (3.05–5.24)	2.84	6
Chen Q. et al.	2018	2018 5551,807	61,129	61,129 5490,678 40.5	40.5	50.8	Sex and age	Register- based	SWE	18–64	ICD-9 ICD-10	17.08 (16.07–18.14)	16.31 (16.21–16.42)	1.05 (0.99–1.11)	2.41	6
Xu et al.	2020	52,821	1642	51,179	45.5	48.6	Age, sex, race/ethnicity, education, family income level, alcohol drinking, smoking, and physical activity, and BMI	Questionnaire	SU	20-79	Retrospective diagnosis	70.04 (58.17–83.47)	88.24 (85.8-90.73)	0.78 (0.66–0.95)	1.54	∞
Note. Inci	dence ris	sk calculated	l per 1000	) population	units. US	= United	Note. Incidence risk calculated per 1000 population units. US = United States, TW = Taiwan, SWE = Sweden, IRR = incidence rate ratio, aOR = Adjusted odds ratio.	= Sweden, IRR :	= incidence	e rate ratio	, aOR = Adjusted	l odds ratio.				

indicating a significant, medium pooled association between ADHD and T2D (Cohen's d = 0.46, Common Language Effect Size [CLES] = 62.74%), t(3) = 6.03, p = 0.009. Fig. 2 displays the results of the metaanalysis. According to the Q-test, there was a significant heterogeneity, i.e., the variance in the ORs was larger than what can be attributed to sampling error ( $\chi^2$  (3) = 10.84, p = 0.013,  $\hat{\tau} = 0.23$ , 95 % CI  $[0.05-1.04]; I^2 = 77.69\%, 95\%$  CI [15.99-98.61]). After examining the studentized residuals, one study (Xu et al., 2021) showed values greater than  $\pm$  2.49, suggesting that it may be deemed as a potential outlier in the context of this model. Based on the Cook's distances, none of the studies seemed to be overly influential. Sensitivity analysis using the leave-one-out method showed that the pooled effect remained largely unchanged (OR range 2.16-2.43), however, after removing Xu et al. (2021) study, between-study variation substantially decreased ( $\chi^2(2)$ ) = 1.24, p = 0.54;  $I^2 = 0\%$ ) and the pooled effect size increased (OR=2.43). Power analyses showed excellent statistical power (0.99) for both the summary of the effect sizes and heterogeneity tests. Results from the meta-regression did not provide evidence for presence of any influence from NOS scores, total sample size, or year of publication on the effect sizes (BF\_{10} = 0.676, BF\_{10} = 0.845, and BF\_{10} = 0.623, respectively). The trim-and-fill method suggested one potential missing study, however, including this possible missing study did not significantly change the association or its significance (OR = 2.16, 95 % CI [1.69-2.75]). Moreover, the selection model using a half-normal function showed a nonsignificant publication bias, likelihood ratio test  $\chi^2$ = 0.09, p = 0.756.

#### 3.2. Population-based study

The cohort comprised 4,216,216 individuals (3,226,030 individuals nested within 1,306,841 families with at least 2 full-siblings, [Median = 2, range = 2–16]) of whom 2,158,775 (51%) were men and 34,715 (0.82%) had an ADHD diagnosis (see Table 2). We followed individuals for a total of 52,873,533 person-years (Md = 13, SD = 1.85) with a median age at start of follow-up of 38.13 years (IQR = 28–50). The ageand sex-adjusted incidence rate of T2D in individuals with ADHD was 51.1, 95% CI (47.9–54.5) cases per 10,000 person-years and 23.0, 95% CI (22.8–23.1) in individuals without ADHD. Supplementary Fig. S1 displays cumulative hazard functions for the overall cohort and each age category.

In the baseline model, individuals with ADHD were at increased risk of developing T2D, HR = 2.35, 95% CI (2.14–2.58). The association between T2D and ADHD attenuated after further adjusting for education and psychiatric comorbidity, HR = 1.21 (1.10–1.33). The HR associated with ADHD further decreased when also adjusting for use of antipsychotic drugs, HR = 1.13, 95% CI (1.03–1.25), with a similar risk

between males and females (HR = 1.14, 95% CI [1.01–1.29], HR = 1.14, 95% CI [0.98–1.33], respectively). BIC indicated that the fully adjusted model provided the best model fit to the data and the concordance index showed a strong discrimination power. Baseline sibling comparisons showed a similar magnitude of associations to the general population (i. e., between-individuals analyses) for developing T2D in individuals with ADHD compared to their undiagnosed full siblings, HR = 2.22 (1.78–2.76).

When separately adjusting each set of factors (education, psychiatric comorbidities, psychiatric comorbidities and antipsychotic use, and cardiometabolic risk factors), we observed that among those risk factors, psychiatric comorbidities attenuated the most the relationship between ADHD and T2D from 2.08 to 1.21 (see Table 3). Given that psychiatric comorbidities showed a substantial impact on the observed association, we explored which specific psychiatric comorbidities contributed the most by adjusting the baseline model for each psychiatric comorbidity in separate models. Amongst the psychiatric comorbidities, substance use disorder (SUD) seemed to be the main driver of the association between ADHD and T2D, followed by anxiety and depression with a similar magnitude (Table 3). Neither bipolar disorder or schizophrenia seemed to have a strong influence on the association between ADHD and T2D.

Sensitivity analyses showed that when analyzing ADHD as a timefixed variable (i.e., individuals who were diagnosed with ADHD during follow up were assumed to have ADHD since the start of follow up), result remained unchanged in terms of both estimates and significance (Supplementary Table S3). Moreover, when assessing the relationship between ADHD and T2D at different ages, the magnitude of the association seemed to decrease with age and was no longer significant in the oldest individuals (51–73 years), range of HR = 1.24–3.84. For the complete estimates see Supplementary Table S4.

#### 4. Discussion

In this study, we systematically reviewed and meta-analyzed all studies that met our inclusion criteria assessing the relationship between ADHD and T2D. In addition, we performed a nationwide populationbased register-linkage study and sibling comparisons to further clarify the role of potential mediators and unmeasured familial factors. The results from the meta-analysis of 5,738,287 individuals (103,022 with ADHD) from 4 studies showed that individuals with ADHD had more than a two-fold increased risk of developing T2D compared to those without ADHD, thus showing a medium-sized association between ADHD and T2D (Cohen's d = 0.46, CLES = 62.68 %).

Similarly, results from the population-based study revealed a more than 2-fold increased risk of T2D (HR = 2.35, 95 % CI [2.14-2.58]), with a similar risk found in males and females. Adjustments for measured and

	N T2I	D / total			
Author(s)	ADHD	No ADHD	Adjusted Odds ratio (aOR), (95% Cl)	Weight (%)	aOR (95% Cl)
Chen HJ. et al.	36/4302	64/21510		18.8	2.75 (1.82 to 4.16)
Chen MH. et al.	158/35949	79/71898		22.3	2.84 (2.03 to 3.97)
Chen Q. et al.	1044/61129	89565/5490678	<b></b>	34.0	2.41 (2.27 to 2.56)
Xu et al.	115/1642	4516/51179		24.9	1.54 (1.16 to 2.04)
Total (95% CI) Test for heterogeneity: r Test for overall effect: Z	$t^2=0.05; \chi^2=10.84, d$	94224/5635265 f=3, P=0.01; I <sup>2</sup> =78% 1	2 3 4	100.0	2.29 (1.48 to 3.55)

Fig. 2. Forest plot of the observed adjusted odds ratios (OR) and the estimate of the random-effects model for the association of ADHD with T2D with 95 % confidence interval, weight, heterogeneity, and overall effect. *Note.* T2D=Type 2 diabetes, CI=confidence interval, aOR=adjusted odds ratio.

#### Table 2

Descriptive statistics of the study population stratified by individuals with and without an ADHD diagnosis.

Variable	Overall,	Without ADHD,	With ADHD,	p-
Vallable	$N = 4216,216^{a}$	$N = 4181,501^{a}$	$N = 34,715^{a}$	value <sup>b</sup>
Demographics				
Sex				< 0.001
Male	2158,775	2139,484	19,291	0.001
indic	(51%)	(51%)	(56%)	
Female	2057,441	2042,017	15,424	
	(49%)	(49%)	(44%)	
Education				< 0.001
Primary and	592,902 (14%)	583,916 (14%)	8986 (26%)	
lower				
secondary				
Upper	1909,431	1892,599	16,832	
secondary	(45%)	(45%)	(48%)	
Postsecondary	1471,391	1463,982	7409 (21%)	
	(35%)	(35%)		
Postgraduate	46,269 (1.1%)	46,136 (1.1%)	133 (0.4%)	
N/A	196,223	194,868 (4.7%)	1355 (3.9%)	
	(4.7%)			
Age at follow up	38 (12)	38 (12)	38 (10)	0.6
Person-years	52,873,533	52,421,995	451,538	0.001
Median Person-	13.00 (2.01)	13.00 (1.85)	13.0 (1.50)	< 0.001
years				
Physical conditions				
T2D	119,055	118,625 (2.8%)	430 (1.2%)	< 0.001
	(2.8%)			
Anxiety	182,466	167,494 (4.0%)	14,972	< 0.001
	(4.3%)		(43%)	
Autism	12,896 (0.3%)	8887 (0.2%)	4009 (12%)	< 0.001
Bipolar disorder	41,275 (1.0%)	36,218 (0.9%)	5057 (15%)	< 0.001
Conduct disorder	1695 (<0.1%)	1276 (<0.1%)	419 (1.2%)	< 0.001
Depression	227,537	212,111 (5.1%)	15,426	< 0.001
-	(5.4%)		(44%)	
Eating disorders	10,859 (0.3%)	9857 (0.2%)	1002 (2.9%)	< 0.001
Hyperlipidemia	141,748	141,214 (3.4%)	534 (1.5%)	< 0.001
	(3.4%)			
Intellectual	18,361 (0.4%)	17,255 (0.4%)	1106 (3.2%)	< 0.001
disability				
Obesity	112,832	110,590 (2.6%)	2242 (6.5%)	< 0.001
	(2.7%)			
Personality	57,251 (1.4%)	49,678 (1.2%)	7573 (22%)	< 0.001
disorders	24 220 (0 60/)	22 402 (0 60/)	007 (0 40/)	< 0.001
Schizophrenia Sleep disorders	24,239 (0.6%)	23,402 (0.6%) 136,974 (3.3%)	837 (2.4%) 4187 (12%)	$< 0.001 \\ < 0.001$
sleep disorders	141,161 (3.3%)	130,974 (3.3%)	4107 (12%)	< 0.001
Substance use	195,326	181,740 (4.3%)	13,586	< 0.001
disorder	(4.6%)	101,7 10 (1.070)	(39%)	0.001
Antipsychotic	131,493	119,367 (2.8%)	12,126	< 0.001
medications	(3.1%)	,007 (2.070)	(34%)	0.001
medications	(0.170)		(01/0)	

N/A=Not available

<sup>a</sup> n (%); Median (SD)

 $^{\rm b}\,$  Pearson's Chi-squared test; Wilcoxon rank sum test

unmeasured familial factors suggested a statistically significant association between ADHD and T2D, and that the observed association, while remaining significant, was largely explained by psychiatric comorbidities, in particular SUD, depression, and anxiety. Further, unmeasured familial factors (i.e., genetic and environmental) shared between siblings appeared to be of limited importance, as evidenced by a recent Swedish co-aggregation study (Du Rietz et al., 2021). Our findings contribute to the available literature in three important ways. First, the increased risk for development of T2D observed in individual studies may largely be explained by psychiatric comorbidities that may mediate this relationship (e.g., ADHD increases the risk of SUD, which in turn increases the risk of T2D). This is a novel finding given that previous studies included in the meta-analyses did not include psychiatric comorbidities as covariates in their analyses. Second, cardiovascular risk factors and antipsychotic medications also had an impact on the relationship between ADHD and T2D, but with a smaller effect size

#### Table 3

Results from the Cox regression model displaying the association between ADHD and T2D.

Model	T2D	Concordance	BIC
	120	concortantee	210
Sequential adjustment			
Baseline (sex and birth year) <sup>1</sup>	2.35 (2.14–2.58)	0.571	3278437.83
Psychiatric disorders and cardiometabolic risk factors <sup>2</sup>	1.21 (1.10–1.33)	0.678	3232592.64
Psychiatric disorders and cardiometabolic risk	1.13 (1.03–1.25)	0.679	3232135.95
factors antipsychotics <sup>3</sup> Fully adjusted by sex			
Males	1.14 (1.01–1.29)	0.643	1966939.52
Females	1.14(0.98-1.33)	0.673	1108204.72
Individual adjustment <sup>4</sup>	1.14 (0.96–1.55)	0.073	1106204.72
Education	2.08 (1.89-2.29)	0.631	3258690.00
Psychiatric disorders	1.32 (1.20–1.45)	0.597	3271714.88
Psychiatric disorders/	1.21 (1.09–1.33)	0.600	3270961.37
antipsychotics	1.21 (1.09-1.33)	0.000	32/0901.3/
Cardiometabolic risk	1.91 (1.74-2.11)	0.627	3253948.07
factors	,	,	
Psychiatric disorder <sup>4</sup>			
Anxiety	1.78 (1.62-1.96)	0.581	3276279.81
Depression	1.78 (1.61-1.96)	0.584	3275674.76
SUD	1.73 (1.57-1.91)	0.585	3274990.69
Bipolar	2.05 (1.86-2.26)	0.575	3277631.86
Schizophrenia	2.26 (2.05-2.48)	0.576	3276909.77
Sibling analyses			
Baseline (sex and birth year) <sup>1</sup>	2.22 (1.78–2.76)	0.579	99694.10

*Note.* <sup>1</sup>Model adjusted for birth year and sex. <sup>2</sup>Model adjusted for birth year, sex, education, anxiety, hyperlipidemia, obesity, sleep disorders, depression, schizophrenia, bipolar disorder, and substance use. <sup>3</sup>Model adjusted for birth year, sex, education, anxiety, hyperlipidemia, obesity, sleep disorders, depression, schizophrenia, bipolar disorder, substance use, and antipsychotic use. <sup>4</sup>Model adjusted for birth year and sex. Bolded estimates display FDR-adjusted *p*-values < 0.05.

compared to psychiatric comorbidities. An extensive body of literature highlighted an association between ADHD and obesity (Cortese et al., 2016; Leppert et al., 2020), however, our results showed a small effect and that the ADHD-T2D relationship remained significant after adjustments for obesity and other well-established cardiovascular risk factors (Garcia-Argibay et al., 2022). Third, our results indicated that unmeasured familial factors seem to have minimal impact on this relationship. This finding is not necessarily inconsistent with evidence from a recent large-scale genome-wide association study (GWAS) on ADHD suggesting a significant genetic correlation between ADHD and T2D ( $r_g = 0.18$ ; Demontis et al., 2019). It is important to highlight that the observed genetic correlation between ADHD and T2D was weak and that the sibling comparison design, used in the current study, only accounts for genetic factors shared by full siblings, which is on average 50 % of their segregating genes. One plausible explanation for the findings of the current study is that ADHD is an important risk factor for T2D and that psychiatric problems (e.g., SUD) that typically emerge after ADHD may mediate this association.

There are several possible explanations for the mediating effect of SUD, depression, and anxiety with T2D. These include behavioral factors, such as unhealthy behaviors (i.e., alcohol consumption and smoking) and dietary habits (i.e., skipping meals/overeating and skipping physical activity) and neurobiological abnormalities. For instance, SUD could potentially worsen the symptoms of ADHD (Perugi et al., 2019), engaging in behaviors that increase the risk of developing T2D, such as poor diet and lack of exercise. Individuals with anxiety disorders (e.g., generalized anxiety disorder) may experiment physical symptoms, such as increased heart rate and rapid breathing (Stein and Sareen, 2015), which could make it more difficult for a person to engage in healthy behaviors that can prevent the development of T2D or engage in

unhealthy eating patterns that increase the likelihood of T2D. A similar pattern could arise in depressed individuals, by which symptoms of depression (e.g., anhedonia, tiredness, and lack of energy (American Psychiatric Association, 2013)) may also lead to unhealthy behaviors including eating disorders or lack of physical exercise. In terms of neurobiological abnormalities, one hypothesis could be a dysregulated hypothalamic-pituitary-adrenal (HPA) axis. Both T2D and anxiety are with associated increased activity in the hypothalamic-pituitary-adrenal (HPA) axis and, therefore, increased secretion of cortisol (Joseph and Golden, 2017). For example, it has been shown that cortisol levels are elevated in individuals with T2D (Hackett et al., 2014; Joseph and Golden, 2017; Liu et al., 2005) and people with a history of anxiety disorders (Chaudieu et al., 2008; Mantella et al., 2008). A similar HPA dysregulation can also be seen in SUD (Huizink et al., 2006; Thayer et al., 2006). Further, it is possible that in people with ADHD, depression may contribute to the development of T2D by causing changes in the body that increase the risk of developing the condition. For example, depression has been shown to affect the way the body processes insulin, which is a hormone that plays a key role in regulating blood sugar levels (Kan et al., 2013; Leonard and Wegener, 2020). If the body is not able to process insulin effectively, this can lead to high blood sugar levels and an increased risk of developing T2D. Additionally, depression may also cause changes in other hormones and chemicals in the body that can affect blood sugar levels and increase the risk of developing T2D, such as pro-inflammatory cytokines and glucocorticoids (Kan et al., 2013). However, more research is needed to fully understand the exact mechanisms by which these conditions may mediate the relationship between ADHD and T2D.

An alternative explanation that cannot be ruled out is that a general disease liability may increase the risk of both multiple mental as well as physical diseases (Cortese et al., 2021). Future research is needed to address these potential explanations. Additional research is also needed to explore the potential impact of ADHD medications on the association between ADHD and T2D (Chen et al., 2018a), of particular interest would be guanfacine, a drug that is approved for the treatment of ADHD by the FDA and that has been linked to weight gain and obesity (Galling et al., 2016).

We found that the association between ADHD and T2D varied as function of age, with stronger associations in young adulthood compared to older age groups. This pattern of results might be caused by ADHD misclassification among the older participants, as ADHD is underdiagnosed and undertreated in the oldest individuals in our cohort (Dobrosavljevic et al., 2020). Alternatively, it could suggest that those who are diagnosed at an early age, and thus, potentially more severe cases, may be at higher risk for T2D than less severe cases. However, due to power constraints caused by a low number of ADHD cases with T2D, we were unable to further adjust estimates for age-stratified analyses. Although ADHD and T2D prevalence is higher in men compared to women (Nordström et al., 2016; Willcutt, 2012), the relationship between ADHD and T2D did not vary as a function of sex.

Our results should be interpreted in the context of a number of limitations. First, only four studies investigated the relationship between ADHD and T2D. Although meta-analyses are statistically appropriate even in cases with a small number of studies, results may be unstable and may vary with the inclusions of new studies. This limitation was addressed by fitting a random effects model and performing a leave-oneout sensitivity analysis, which showed that the association was not driven by one particular study. Second, variability between studies in effect sizes in the meta-analysis reflects methodological differences and limitations related to confounder/mediator control, number of factors adjusted for, and small sample sizes. Several studies lacked an in-depth mediator control, and no study performed a sibling design to account for unmeasured confounding. Lastly, a major limitation of the current literature is that most studies (3/4) were registry-based and, thus, ADHD diagnoses might capture only the most severe cases of the ADHD spectrum who sought specialized medical care. This limitation is also evident in our register-based study. Furthermore, the directionality of the ADHD-T2D association remains unclear, as two of the included studies used a cross-sectional design (Chen et al., 2018b; Xu et al., 2021), and the available longitudinal studies focused on how T2D associates with ADHD (Chen et al., 2013; Chen et al., 2018a). However, results from our register-based study, together with epidemiological data about age of onset for ADHD and T2D (Koopman, 2005; Polanczyk et al., 2010) suggest that individuals with ADHD are more likely to develop T2D. Nevertheless, further studies attempting to determine the directional and causal nature of the relationship between ADHD and T2D should employ other methodology such as Mendelian randomization. Moreover, misclassification of ADHD in older ages is likely given that ADHD is underdiagnosed in adults (Dobrosavljevic et al., 2020). In contrast, one study used ADHD diagnoses from self-reports (Xu et al., 2021), which might capture less severe and subclinical ADHD symptoms. In addition, relying on individual reports to determine the presence of ADHD and T2D may introduce the possibility of faulty recollection affecting the study results. However, results from the leave-one-out sensitivity analysis did not show a change in the significance, and the pooled effect increased when removing that study. Lastly, diagnoses of hyperlipidemia, obesity, and sleep disorders were identified from the NPR and therefore, are imperfect measures for cardiovascular risk factors, possibly underestimating effect on T2D. These diagnoses mainly apply to individuals who were referred for treatment due to other reasons or the most severe cases that developed related complications. Due to this low reliability on those diagnoses, we were not able to explore the role of metabolic syndrome in the ADHD-T2D association. We cannot rule out that ADHD influences SUD, that in turn increases the likelihood of metabolic syndrome and ultimately T2D. Future research is needed to further explore the possible effect that ADHD pharmacological treatment has on this association. Another factor that warrants further attention and has not been explored is the severity of ADHD, given that it is plausible that the relationship between ADHD and T2D varies at different levels of severity.

#### 5. Conclusions

This study revealed a significant association between ADHD and T2D that was largely due to psychiatric comorbidities, in particular SUD, depression, and anxiety. Our findings suggest that clinicians need to be aware of the increased risk of developing T2D in individuals with ADHD and that psychiatric comorbidities may be the main driver of this association. Appropriate identification and treatment of these psychiatric comorbidities may reduce the risk for developing T2D in ADHD, together with efforts to intervene on other modifiable T2D risk factors (e.g., unhealthy lifestyle habits and use of antipsychotics, which are common in ADHD), and to devise individual programs to increase physical activity (Quesada et al., 2018). Considering the significant economic burden of ADHD (Garcia-Argibay et al., 2021) and T2D (Zhang and Gregg, 2017), a better understanding of this relationship is essential for targeted interventions or prevention programs with the potential for a positive impact on both public health and the lives of persons living with ADHD.

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#### Role of the Funding source

The funders had no role in the design and conduct of the study; collection, management, analysis, and interpretation of the data; preparation, review, or approval of the manuscript; and decision to submit the manuscript for publication.

#### Contributors

Dr Garcia-Argibay had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis, Concept and design: Garcia-Argibay, Larsson, Statistical analysis: Garcia-Argibay, Acquisition, analysis, or interpretation of data: All authors, Drafting of the manuscript: Garcia-Argibay, Critical revision of the manuscript for important intellectual content: All authors, Supervision: Larsson.

#### **Ethical approval**

The study had ethical approval from the Regional Ethical Review Board in Stockholm, Sweden (Dnr 2013/862–31/5). The requirement for informed consent was waived because the study was register-based and data on the included individuals were deidentified. The investigation conforms to the 1964 Helsinki declaration and its later amendments or comparable ethical standards.

#### **Declaration of Competing Interest**

Henrik Larsson reported receiving grants from Shire/Takeda Pharmaceuticals during the conduct of the study; personal fees from and serving as a speaker for Shire/Takeda Pharmaceuticals and Evolan Pharma AB outside the submitted work; and sponsorship for a conference on attention-deficit/hyperactivity disorder from Shire Pharmaceuticals outside the submitted work. Johan Jendle reported receiving grants from Novo Nordisk during the conduct of the study; fees from and serving as a speaker for Abbott, Boehringer Ingelheim, Eli Lilly, Medtronic, Nordic Infucare, Novo Nordisk, Sanofi, outside the submitted work. Ebba Du Rietz has served as a speaker for Shire Sweden AB outside the submitted work.

J.A.R.Q was on the speakers' bureau and/or acted as consultant for Janssen-Cilag, Novartis, Shire, Takeda, Bial, Shionogi, Sincrolab, Novartis, BMS, Medice, Rubió, Uriach, Technofarma and Raffo in the last 3 years. He also received travel awards (air tickets + hotel) for taking part in psychiatric meetings from Janssen-Cilag, Rubió, Shire, Takeda, Shionogi, Bial and Medice. The Department of Psychiatry chaired by him received unrestricted educational and research support from the following companies in the last 3 years: Janssen-Cilag, Shire, Oryzon, Roche, Psious, and Rubió.

The remaining authors declare having no conflict of interest.

## Appendix A. Supporting information

Supplementary data associated with this article can be found in the online version at doi:10.1016/j.neubiorev.2023.105076.

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